

## 457.212 Statistics for Civil & Environmental Engineers

### In-Class Material: Class 11

#### Useful Distribution Models – Part I (A&T: 3.2)

“Distribution models” are useful because ...

- The probability function is the result of an underlying physical process and can be derived on the basis of certain physically reasonable assumptions.
- The function is the result of some limiting process (e.g. \_\_\_\_\_).
- It is widely known, and the necessary probability and statistical information (including probability tables) are widely available (e.g. probability table of \_\_\_\_\_)

#### 1. Normal distribution

- Best known and most widely used. Also known as \_\_\_\_\_ distribution.
- According to \_\_\_\_\_, the sum of random variables converges to a normal random variable as the number of the variables increases, no matter what distributions the variables are subjected to.
- Completely defined by the \_\_\_\_\_ and the \_\_\_\_\_ of the random variable.

(a) PDF:  $X \sim N(\mu, \sigma)$

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right], \quad -\infty < x < \infty$$

(b) CDF: no closed-form expression available

$$F_X(x) = \int_{-\infty}^x f_X(x) dx, \quad -\infty < x < \infty$$

(c) Parameters:  $\mu, \sigma$

- $\mu$ : \_\_\_\_\_ of the random variable, i.e.  $\mu = \mu_X \equiv E[X]$
- $\sigma$ : \_\_\_\_\_ of the random variable, i.e.  $\sigma = \sigma_X \equiv \{E[(X - \mu_X)^2]\}^{0.5}$

(d) Shape of the PDF plots

- Symmetric around  $x =$  \_\_\_\_\_
- A change in  $\mu_X$  \_\_\_\_\_ the PDF horizontally by the same amount.
- The larger the value of  $\sigma_X$  gets, the more \_\_\_\_\_ the PDF becomes around the central axis.

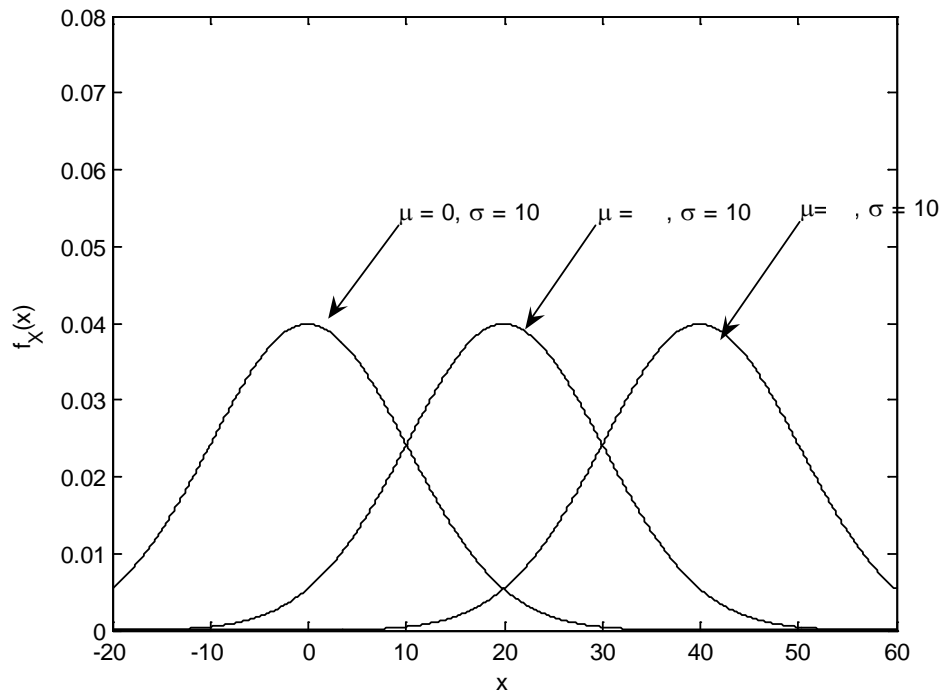


Figure 1. PDF's of normal random variables with different values of  $\mu$

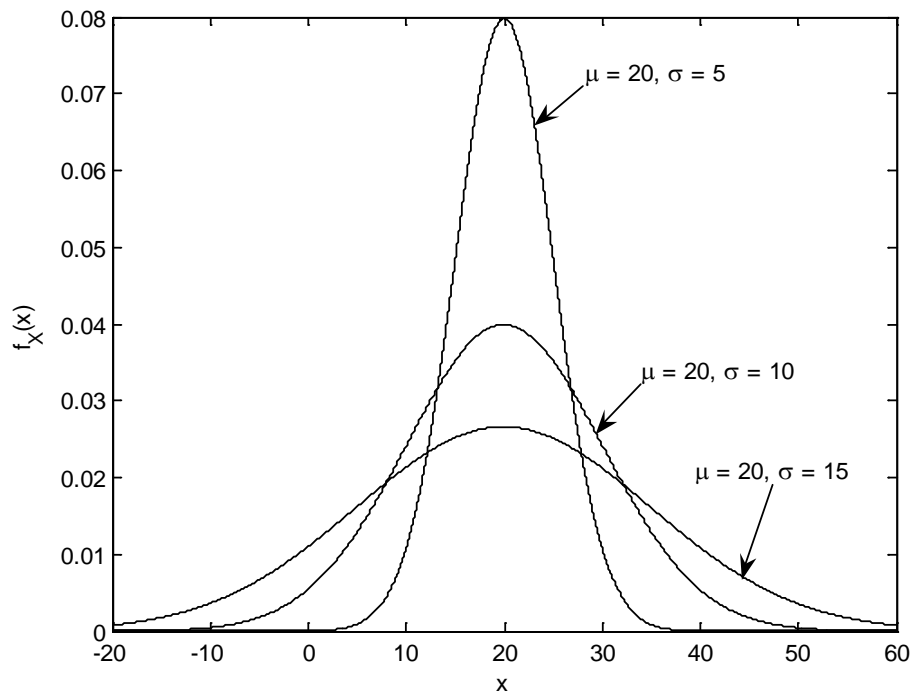


Figure 2. PDF's of normal random variables with different values of  $\sigma$

1a. **Standard** normal distribution

- A special case of the normal distribution:  $\mu_x =$  ,  $\sigma_x =$  .
- The CDF of the standard normal distribution can be used for computing the CDF of any general normal random variable.

(a) PDF:  $U \sim N( , )$

$$\varphi(u) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}u^2\right), \quad -\infty < u < \infty$$

(b) CDF:

$$\Phi(u) = \int_{-\infty}^u \varphi(u) du, \quad -\infty < u < \infty$$

→ no closed-form expression available, but the table of the standard normal CDF  $\Phi(\cdot)$  can be found in books or computer software (e.g. See Appendix A of A&T)

(c) Inverse CDF of standard normal distribution:  $\Phi^{-1}(\cdot)$

$$\Phi(u_p) = p \quad \Leftrightarrow \quad u_p = \Phi^{-1}(p)$$

(d) Symmetry around  $u =$  :

$$\begin{aligned} \Phi(-u) &= 1 - \Phi(u) \\ u_{1-p} &= -u_p \end{aligned}$$

→ The table of the standard normal CDF is often provided for positive  $u$  values only, but using the symmetry one can find the CDF for negative values as well.

(e) One can compute the CDF of a general normal random variable  $X \sim N(\mu, \sigma)$  by use of the CDF of the standard normal random variable  $U \sim N(0,1)$  as follows.

$$\begin{aligned} F_X(a) &= P(X \leq a) \\ &= \int_{-\infty}^a \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right] dx \\ &= \int_{-\infty}^{\left(\frac{a-\mu}{\sigma}\right)} \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}u^2\right) \sigma du \\ &= \Phi\left(\frac{\quad}{\quad}\right) \end{aligned}$$

Hence,  $P(a < X \leq b) = F_X(\quad) - F_X(\quad) = \Phi\left(\frac{\quad}{\quad}\right) - \Phi\left(\frac{\quad}{\quad}\right)$

**Example 1:** Given a standard normal distribution, find the area under the curve that lies

(a) to the right of  $u = 1.84$

(b) between  $u = -1.97$  and  $u = 0.86$

**Example 2:** The drainage demand during a storm (in mgd: million gallons/day):  
 $X \sim N(1.2, 0.4)$ . The maximum drain capacity is 1.5 mgd.

(a) Probability of flooding?

(b) Probability that the drainage demand during a storm will be between 1.0 and 1.6 mgd?

(c) The 90-percentile drainage demand?

2. **Lognormal** distribution

- Closely related to the \_\_\_\_\_ distribution.
- Defined for \_\_\_\_\_ values only.

(a) PDF:  $X \sim LN(\lambda, \zeta)$

$$f_X(x) = \frac{1}{\sqrt{2\pi\zeta x}} \exp\left[-\frac{1}{2}\left(\frac{\ln x - \lambda}{\zeta}\right)^2\right], \quad 0 < x < \infty$$

(b) CDF:

$$F_X(x) = \int_{-\infty}^x f_X(x) dx, \quad 0 < x < \infty$$

→ no closed-form expression available, but can be computed by use of the table of the standard normal CDF  $\Phi(\cdot)$  (as shown below)

(c) Parameters:  $\lambda, \zeta$

- $\lambda$ : mean of \_\_\_\_\_, i.e.  $\lambda = \lambda_X \equiv E[\ln X]$
- $\zeta$ : standard deviation of \_\_\_\_\_, i.e.  $\zeta^2 = \zeta_X^2 = \sigma_{\ln X}^2$

(d) Shape of the PDF plots

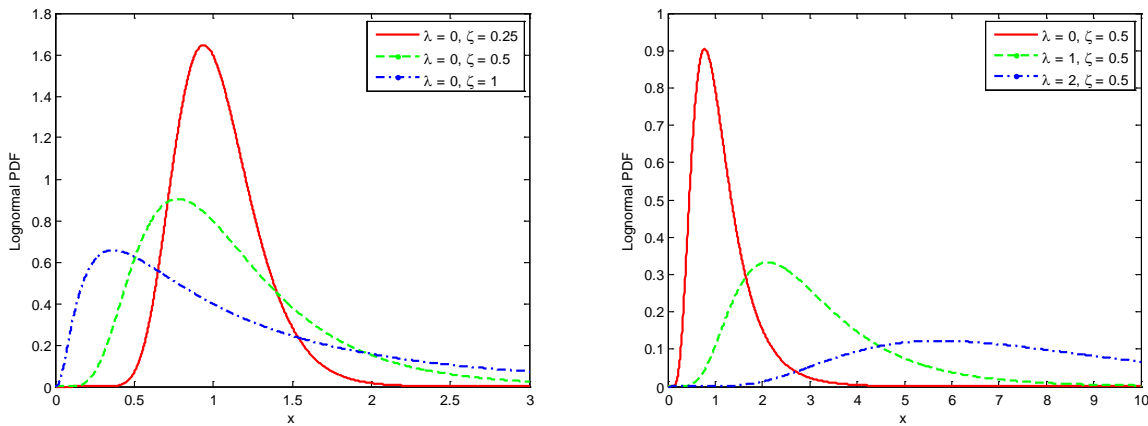


Figure 3. PDF's of lognormal random variables.

(e) Relationship between normal and lognormal distribution:

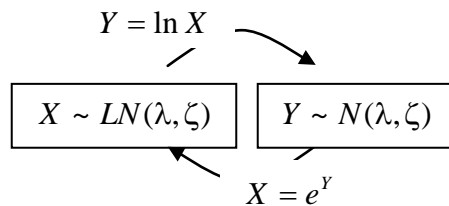
“The logarithm of a \_\_\_\_\_ random variable is a \_\_\_\_\_ random variable.”

$$X \sim LN(\lambda, \zeta) \Rightarrow \ln X \sim N(\lambda, \zeta)$$

(f) Can obtain the CDF of lognormal  $X \sim LN(\lambda, \zeta)$  from the CDF of standard normal:

$$\begin{aligned} F_X(a) &= P(X \leq a) \\ &= P(\ln X \leq \ln a) \quad \text{Since } \ln X \sim N(\lambda, \zeta), \\ &= \Phi\left(\frac{\ln a - \lambda}{\zeta}\right) \end{aligned}$$

(g) “The exponential function of a \_\_\_\_\_ random variable is a \_\_\_\_\_ random variable.”



(h)  $(\lambda, \zeta) \rightarrow (\mu, \delta)$ : Find the mean and c.o.v. from the distribution parameters

$$\begin{aligned} \mu &= E[X] = \exp(\lambda + 0.5\zeta^2) \\ \delta &= \sigma / \mu = \sqrt{\exp(\zeta^2) - 1} \quad (\cong \zeta \text{ for } \zeta \ll 1) \end{aligned}$$

(i)  $(\mu, \delta) \rightarrow (\lambda, \zeta)$ : Find the distribution parameters from the mean and c.o.v.

$$\begin{aligned} \zeta &= \sqrt{\ln(1 + \delta^2)} \quad (\cong \delta \text{ for } \delta \ll 1) \\ \lambda &= \ln \mu - 0.5 \ln(1 + \delta^2) \end{aligned}$$

(j)  $(x_{0.5}) \leftrightarrow (\lambda)$ : Relationship between the median and  $\lambda$

$$\lambda = \ln x_{0.5}, \quad x_{0.5} = e^\lambda$$

(k)  $(\mu, \delta) \rightarrow (x_{0.5})$ : Find the median from the mean and c.o.v.

$$x_{0.5} = \frac{\mu}{\sqrt{1 + \delta^2}}$$

Note:  $x_{0.5} < \mu$  for the lognormal distribution.

**Example 3:** The drainage demand during a storm (in mgd: million gallons/day) is assumed to follow the lognormal distribution with the same mean and standard deviation as Example 1 (mean 1.2, standard deviation 0.4). The maximum drain capacity is 1.5 mgd.

- (a) Distribution parameters, i.e.  $\lambda$  and  $\zeta$ ?
- (b) Probability of the flooding?
- (c) Probability that the drainage demand during a storm will be between 1.0 and 1.6 mgd?
- (d) The 90-percentile drainage demand?

**Example 4:** Consider a bridge whose uncertain capacity against “complete damage” limit-state caused by earthquake events is defined in terms of peak ground acceleration (PGA; unit: g) that the bridge can sustain. Suppose the median of the capacity is 1.03g and the coefficient of variation is 0.50. It is assumed that the capacity follows a lognormal distribution.

- (a) Distribution parameters of the lognormal distribution, i.e.  $\lambda$  and  $\zeta$ ?
- (b) The mean and standard deviation of the uncertain capacity, i.e.  $\mu$  and  $\sigma$ ?
- (c) Suppose the peak ground acceleration from an earthquake event is 0.5g. What is the probability that the structure will exceed “complete damage” limit state?





## Probability Distribution Models in Matlab® Statistics Toolbox

Full Name	Short	Parameters	Probability Density/Mass Function	Mean	Variance
Binomial	<i>binom</i>	$0 < p < 1$ $n$ integer	$\binom{n}{x} p^x (1-p)^{(n-x)}, \quad x = 0, 1, \dots, n$	$np$	$np(1-p)$
Geometric	<i>geom</i>	$0 < p < 1$	$p(1-p)^x, \quad x = 0, 1, 2, \dots$	$(1-p)/p$	$(1-p)/p^2$
Hypergeometric	<i>hyge</i>	$0 < K, N \leq M$ $K, N, M$ integers	$\binom{K}{x} \binom{M-K}{N-x} \binom{M}{N}^{-1}, \quad K+N-M \leq x \leq K$	$\frac{NK}{M}$	$N \frac{K}{M} \frac{M-K}{M} \frac{M-N}{M-1}$
Negative Binomial	<i>nbino</i>	$0 < p < 1$ $r$ integer	$\binom{r+x-1}{x} p^r (1-p)^x, \quad x = 0, 1, \dots$	$r(1-p)/p$	$r(1-p)/p^2$
Poisson	<i>poiss</i>	$0 < \lambda$	$\frac{\lambda^x}{x!} e^{-\lambda}, \quad x = 0, 1, \dots$	$\lambda$	$\lambda$
Beta	<i>betad</i>	$0 < a, b$	$B(a,b)^{-1} x^{a-1} (1-x)^{b-1}, \quad 0 \leq x \leq 1$	$a/(a+b)$	$ab/(a+b+1)/(a+b)^2$
Chisquare	<i>chi2</i>	$0 < v$	$x^{(v-2)/2} e^{-x/2} 2^{-v/2} \Gamma(v/2)^{-1}, \quad 0 < x$	$v$	$2v$
Exponential	<i>expd</i>	$0 < \mu$	$\mu^{-1} e^{-x/\mu}, \quad 0 < x$	$\mu$	$\mu^2$
F	<i>f</i>	$0 < v_1, v_2$	$\frac{\Gamma((v_1+v_2)/2) (v_1/v_2)^{v_1/2} x^{v_1/2-1}}{\Gamma(v_1/2) \Gamma(v_2/2) [1+(v_1/v_2)x]^{(v_1+v_2)/2}}, \quad 0 < x$	$v_2/(v_2-2)$	$\frac{2v_2^2(v_1+v_2-2)}{v_1(v_2-2)^2(v_2-4)}$
Gamma	<i>gamd</i>	$0 < a, b$	$b^{-a} \Gamma(a)^{-1} x^{a-1} e^{-x/b}, \quad 0 < x$	$ab$	$ab^2$
Lognormal	<i>lognd</i>	$\lambda, 0 < \zeta$	$x^{-1} \zeta^{-1} (2\pi)^{-1/2} \exp[-(\ln x - \lambda)^2 / 2\zeta^2], \quad 0 < x$	$e^{(\lambda+0.5\zeta^2)}$	$e^{(2\lambda+2\zeta^2)} - e^{(2\lambda+\zeta^2)}$
Normal	<i>normd</i>	$\mu, 0 < \sigma$	$\sigma^{-1} (2\pi)^{-1/2} \exp[-(x-\mu)^2 / 2\sigma^2]$	$\mu$	$\sigma^2$
Rayleigh	<i>rayld</i>	$0 < b$	$x b^{-2} \exp(-x^2/2b^2), \quad 0 < x$	$b\sqrt{\pi/2}$	$(4-\pi)b^2/2$
T	<i>td</i>	$0 < v$	$(v\pi)^{-1/2} \Gamma((v+1)/2) \Gamma(v/2)^{-1} (1+x^2/v)^{-(v+1)/2}$	0	$v/(v-2)$
Uniform	<i>unifd</i>	$a < b$	$(b-a)^{-1}, \quad a \leq x \leq b$	$(a+b)/2$	$(b-a)^2/12$
Weibull	<i>weibd</i>	$0 < a, b$	$abx^{b-1} e^{-ax^b}, \quad 0 < x$	$a^{-1/b} \Gamma(1+b^{-1})$	$a^{-2/b} [\Gamma(1+2b^{-1}) - \Gamma^2(1+b^{-1})]$

Use *shortnamepdf()* to compute the probability density/mass function; *shortnamecdf()* to compute cumulative distribution function; *shortnamefit()* to estimate parameters from data; *shortnamernd()* to generate random numbers; *shortnamestat()* to compute mean and variance for specified parameters; and *shortnameinv()* to compute the inverse cumulative probability. Use Matlab® help to learn more about these commands.