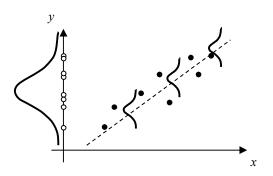
# 457.212 Statistics for Civil & Environmental Engineers In-Class Material: Class 25 Regression Analysis (A&T: 8.2-8.4, 8.7)

Given: Sample data set  $\{(x_1, y_1), (x_2, y_2), ..., (x_n, y_n)\}$ Question: The functional relation between two random variables X and Y? Y = f(X)→ "Regression" Analysis

# 1. Regression & Conditional Mean



(a) Marginal and conditional standard deviation of Y:

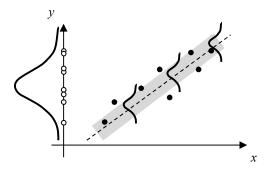
$$\sigma_{Y} = \sigma_{Y|x}$$

- (b) Marginal and conditional mean of Y:
  - E[Y] = constant.E[Y | x] = f(x)
  - $\rightarrow$  Conditional mean predicts the outcome of Y more accurately (i.e. smaller variation). → Regression analysis aims at finding the functional relationship for the conditional mean to describe the hidden relation between X and Y.
- (c) Linear vs. nonlinear regression v  $x_2$ x  $x_1$



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#### 2. Single Linear Regression with Constant (Conditional) Variance



(a) Assumption: the conditional mean is a linear function of x and the conditional variance is constant, i.e.

$$E[Y | x] = \alpha + \beta x$$
 and  $\sigma_{Y|x}^2 = const.$ 

"Linear regression of Y on X"

(b) Estimation of  $\alpha$  and  $\beta$ 

"Best" estimates on  $\alpha$  and  $\beta$ :  $\hat{\alpha}$  and  $\hat{\beta}$  ~ the values minimizing the sum of squared errors between the prediction by the linear relationship ( $y'_i = \alpha + \beta x_i$ ) and the given data point  $y_i$  (least square estimators)

Sum of Squared Errors (SSE):

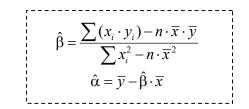
$$\Delta^{2} = \sum_{i=1}^{n} (y_{i} - y'_{i})^{2}$$
$$= \sum_{i=1}^{n} (y_{i} - \alpha - \beta x_{i})^{2}$$

**Note**: The same weight is given to each data point because the conditional variance is assumed to be constant.

Find  $\alpha$  and  $\beta$  that minimize SSE  $\rightarrow$  Solve the following equations for  $\alpha$  and  $\beta$ :

$$\frac{\partial \Delta^2}{\partial \alpha} = 2\sum_{i=1}^n (y_i - \alpha - \beta x_i)(-1) =$$
$$\frac{\partial \Delta^2}{\partial \beta} = 2\sum_{i=1}^n (y_i - \alpha - \beta x_i)(-x_i) =$$

As a result,



Need: 
$$\sum x_i y_i$$
,  $\sum x_i$ ,  $\sum y_i$  and  $\sum x_i^2$ 

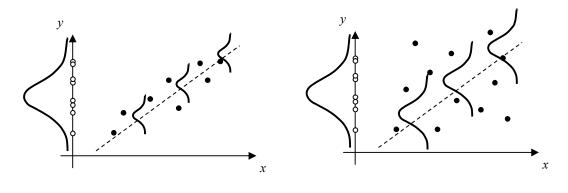
(c)  $\sigma_{Y|x}^2$ ?

Estimated as

$$s_{Y|x}^{2} = \frac{1}{n-2} \sum_{i=1}^{n} (y_{i} - y_{i}')^{2}$$
$$= \frac{\Delta^{2}}{n-2}$$

(d) Reduction of variance: from marginal  $\sigma_{Y}^{2}(s_{Y}^{2})$  to conditional variance  $\sigma_{Y|x}^{2}(s_{Y|x}^{2})$ ?

 $\rightarrow$  A measure of the strength of the linear relationship

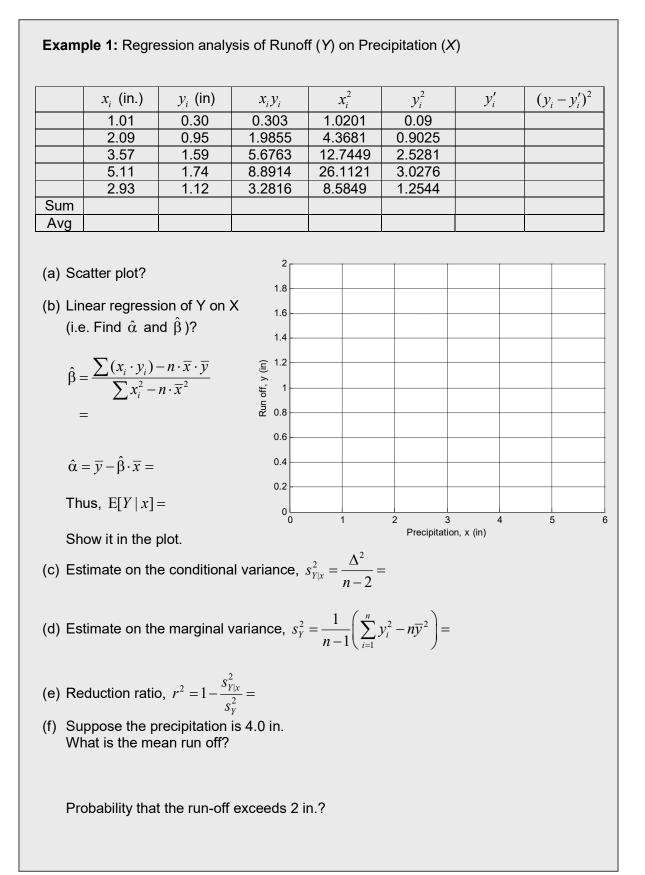


 $r^{2} = \frac{s_{Y}^{2} - s_{Y|x}^{2}}{s_{Y}^{2}} = 1 - \frac{s_{Y|x}^{2}}{s_{Y}^{2}}$ 

 $r^2 \cong 0$ : No reduction (weak linear relationship)

 $r^2 \cong 1$ : Large reduction (strong linear relationship)

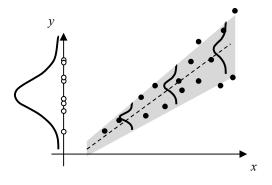
**Note:**  $r^2 \cong \rho_{XY}$  as  $n \to \infty$ 



#### 4

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### 3. Single Linear Regression with Non-constant Variance



(a) Assumption: the conditional mean is a linear function of x and the conditional variance is a function of x, i.e.

$$E[Y | x] = \alpha + \beta x$$
  

$$\sigma_{Y|x} = \sigma \cdot g(x) \quad \text{(Thus, } \sigma_{Y|x}^2 = \sigma^2 g^2(x) \text{)}$$

- e.g.  $\sigma_{Y|x} = \sigma x$  (linearly increasing)  $\sigma_{Y|x} = \sigma x^2$  (quadratically)
- (a) Estimation of  $\alpha$  and  $\beta$

The same as regression with constant variance except that the errors are given nonequal weights.

Sum of Weighted Squared Errors (SWSE):

$$\Delta^2 = \sum_{i=1}^n w'_i \cdot (y_i - y'_i)^2$$
$$= \sum_{i=1}^n w'_i \cdot (y_i - \alpha - \beta x_i)^2$$

Note: Give more weights to the data points that require more accurate fitting.

$$w_i' \equiv \frac{1}{\sigma_{Y|x}^2} = \frac{1}{\sigma^2 g^2(x)}$$

Find  $\alpha$  and  $\beta$  that minimize SWSE  $\rightarrow$  Solve the following equations for  $\alpha$  and  $\beta$ :

$$\frac{\partial \Delta^2}{\partial \alpha} = 0$$
 and  $\frac{\partial \Delta^2}{\partial \beta} = 0$ 

As a result,

$$\hat{\beta} = \frac{(\Sigma w_i)(\Sigma w_i x_i y_i) - (\Sigma w_i y_i)(\Sigma w_i x_i)}{(\Sigma w_i)(\Sigma w_i x_i^2) - (\Sigma w_i x_i)^2}$$
$$\hat{\alpha} = \frac{(\Sigma w_i y_i) - \hat{\beta}(\Sigma w_i x_i)}{\Sigma w_i}$$

where 
$$w_i = \sigma^2 w'_i = \frac{1}{g^2(x_i)}$$

Need: 
$$\sum w_i x_i y_i$$
,  $\sum w_i x_i$ ,  $\sum w_i y_i$ ,  $\sum w_i x_i^2$  and  $\sum w_i$ 

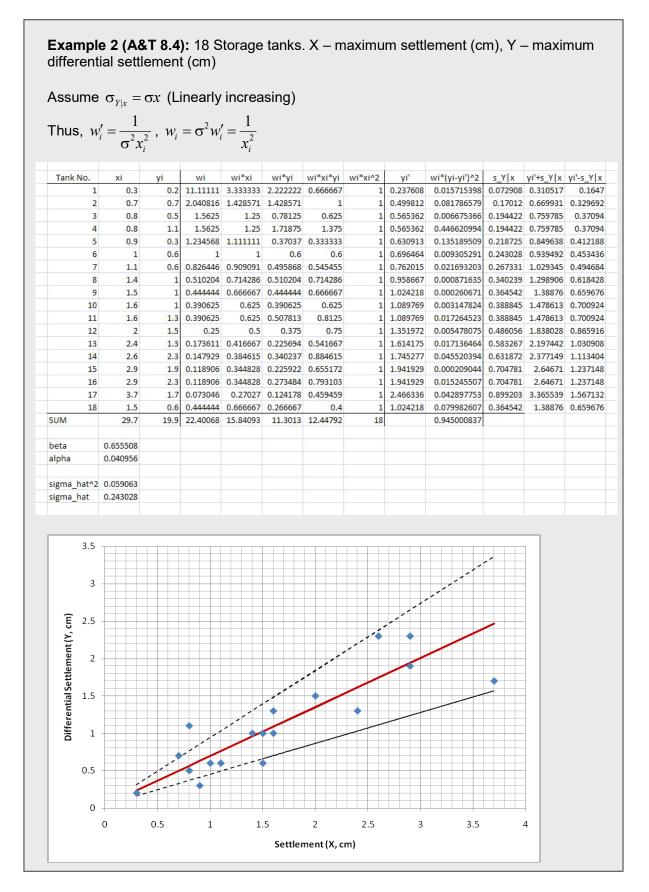
(c) 
$$\sigma_{Y|x}^2$$
?

First, an unbiased estimate of  $\sigma^{2}$  (not  $\sigma^{2}_{\textit{Y}|\textit{x}}$ ) is

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^n w_i (y_i - y'_i)^2}{n-2}$$

Then, the conditional variance is estimated as

$$s_{Y|x}^2 = \hat{\sigma}^2 g^2(x)$$
$$= \frac{\sum w_i (y_i - \hat{\alpha} - \hat{\beta} x_i)^2}{n - 2} g^2(x)$$



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# 4. Multiple Linear Regression

"Linear regression of  $\boldsymbol{Y}$  on  $X_1,...,~X_m$  "

(a) Define  $\Delta^2$  by assuming  $\sigma_{Y|x}^2 = \sigma^2$  (constant) or  $\sigma_{Y|x}^2 = \sigma^2 g^2(x_1, ..., x_m)$  (non-constant) (b) Find

$$E[Y | x_1, ..., x_m] = \beta_0 + \beta_1 x_1 + ... + \beta_m x_m$$

(c) Estimate  $\hat{\beta}_0, \hat{\beta}_1, ..., \hat{\beta}_m$  by solving

$$\frac{\partial \Delta^2}{\partial \beta_0} = \frac{\partial \Delta^2}{\partial \beta_1} = \dots = \frac{\partial \Delta^2}{\partial \beta_m} = 0$$

(d) 
$$s_{Y|x_1,...,x_m}^2 = \frac{\Delta^2}{n-m-1}$$

(**Note**: m = 1 for single linear regression)

### 5. Nonlinear Regression & Applications of Regression Analysis (Read A&T 8.6-8.7)

# 6. Correlation Analysis

(a) (True or theoretical) correlation coefficient

$$\rho_{XY} = \frac{Cov[X,Y]}{\sigma_X \sigma_Y} = \frac{E[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X \sigma_Y} = \frac{E[XY] - \mu_X \mu_Y}{\sigma_X \sigma_Y}$$

(b) Unbiased estimator of  $\rho_{\it XY},~\hat{\rho}$ 

$$\hat{\rho} = \frac{1}{n-1} \frac{\sum_{i=1}^{n} x_i y_i - n\overline{X}\overline{Y}}{s_x s_y}$$

(c)  $\hat{\rho}$  and  $\hat{\beta}$ 

$$\hat{\rho} = \hat{\rho} \frac{s_X}{s_X} = \frac{\Sigma(x_i - \overline{X})(y_i - \overline{Y})}{(n-1)s_X^2} \frac{s_X}{s_Y} = \frac{\Sigma(x_i - \overline{X})(y_i - \overline{Y})}{\Sigma(x_i - \overline{X})^2} \frac{s_X}{s_Y} = \hat{\beta} \frac{s_X}{s_Y}$$

(d)  $\hat{\rho}^2$  and  $r^2 = 1 - \frac{s_{Y|x}^2}{s_Y^2} / \frac{s_Y^2}{s_Y^2}$  $\hat{\rho}^2 = 1 - \frac{n-2}{n-1} \frac{s_{Y|x}^2}{s_Y^2}$ . As  $n \to \infty$ ,  $\hat{\rho}^2 \to 1 - \frac{s_{Y|x}^2}{s_Y^2} = r^2$