

Independence Test for Random Numbers

- In testing for independence, the hypotheses are as follows:

$$H_0: R_i \sim \text{independent}$$

$$H_1: R_i \not\sim \text{independent}$$

- The null hypothesis, H_0 , reads that the numbers are independent. Failure to reject the null hypothesis means that no evidence of dependence has been detected on the basis of this test. This does not imply that further testing of the generator for independence is unnecessary.

Independence Test for Random Numbers

1. Gap test

Counts the number of digits that appear between repetitions of a particular digit and then uses the Kolmogorov-Smirnov test to compare with the expected number of gaps.

2. Runs test

Tests the runs up and down or the runs above and below the mean by comparing the actual values with expected values. The statistic for comparison is the t -statistic.

Independence Test for Random Numbers (cont.)

3. Poker test

Treats numbers grouped together as a poker hand (= each playing card) Then the hands obtained are compared with what is expected using the chi-square test.

4. Autocorrelation test

Tests the correlation between numbers and compares the sample correlation with the expected correlation of zero.

Gap Test for Random Numbers

- The Gap Test measures the number of digits between successive occurrences of the same digit.

(Example) length of gaps associated with the digit 3.

4, 1, 3, 5, 1, 7, 2, 8, 2, 0, 7, 9, 1, 3, 5, 2, 7, 9, 4, 1, 6, 3
3, 9, 6, 3, 4, 8, 2, 3, 1, 9, 4, 4, 6, 8, 4, 1, 3, 8, 9, 5, 5, 7
3, 9, 5, 9, 8, 5, 3, 2, 2, 3, 7, 4, 7, 0, 3, 6, 3, 5, 9, 9, 5, 5
5, 0, 4, 6, 8, 0, 4, 7, 0, 3, 3, 0, 9, 5, 7, 9, 5, 1, 6, 6, 3, 8
8, 8, 9, 2, 9, 1, 8, 5, 4, 4, 5, 0, 2, 3, 9, 7, 1, 2, 0, 3, 6, 3
(22 x 5 = 110 digits)

Note: eighteen 3's in list

=> 17 gaps, the first gap is of length 10

Gap Test for Random Numbers (cont.)

- We are interested in the frequency of gaps.

$$P(\text{gap of } 10) = P(\text{not } 3) \times \dots \times P(\text{not } 3) \times P(3)$$

note: there are 10 terms of the type $P(\text{not } 3) = 0.9$

$$P(\text{gap size or gap length} = 10) = (0.9)^{10} (0.1)$$

- The theoretical cumulative frequency distribution of the gap size x for randomly ordered digit is given by

$$F(x) = 0.1 \sum_{n=0}^x (0.9)^n = 1 - 0.9^{x+1}$$

for $x = 0, 1, 2, \dots$ (the max. gap size or $108 = 110 - 2$)

Note: The observed frequencies for all digits are compared to the theoretical frequency.

Gap Test for Random Numbers (cont.)

Example.

- Based on the frequency with which gaps occur for the digit 3, analyze the 110 digits above to test whether they are independent. Use $\alpha = 0.05$. The total number of gaps (or the number of samples) is given by the number of digits minus 10, or 100.

$$\begin{aligned} * \text{Total number of digits} &= \sum_{i=0}^9 (\text{the number of } i\text{-th digit}) \\ &= \sum_{i=0}^9 (n_i + 1) \end{aligned}$$

where n_i is the number of gaps for the i -th digit.

Thus, total number of gaps $\sum_{i=0}^9 n_i$
 $= \text{total number of digits} - 10.$

- The number of gaps associated with the various digits are as follows:

Digit	0	1	2	3	4	5	6	7	8	9
# of Gaps	7	8	8	17	10	13	7	8	9	13

(total 100 gaps)

Gap Test for Random Numbers (cont.)

Example (cont.)

$\nwarrow F(x)$

Gap Length	Frequency	Relative Frequency	Cum. Rel. Frequency, $F(x)$	$S_M(x)$	$ F(x) - S_M(x) $
0-3	35	0.35	0.35	0.3439	0.0061
4-7	22	0.22	0.57	0.5695	0.0005
8-11	17	0.17	0.74	0.7176	0.0224
12-15	9	0.09	0.83	0.8147	0.0153
16-19	5	0.05	0.88	0.8784	0.0016
20-23	6	0.06	0.94	0.9202	0.0198
24-27	3	0.03	0.97	0.9497	0.0223
28-31	0	0.00	0.97	0.9657	0.0043
32-35	0	0.00	0.97	0.9775	0.0075
36-39	2	0.02	0.99	0.9852	0.0043
40-43	0	0.00	0.99	0.9903	0.0003
44-47	1	0.01	1.00	0.9936	0.0064

D

(total 100 gaps)

Table 1 Critical values, $CV(\alpha, n)$, of the KS test
with sample size n at the different levels of α .

n	Level of significance (α)					
	0.40	0.20	0.10	0.05	0.04	0.01
5	0.369	0.447	0.509	0.562	0.580	0.667
10	0.268	0.322	0.368	0.409	0.422	0.487
20	0.192	0.232	0.264	0.294	0.304	0.352
30	0.158	0.190	0.217	0.242	0.250	0.290
50	0.123	0.149	0.169	0.189	0.194	0.225
>50	$\frac{0.87}{\sqrt{n}}$	$\frac{1.07}{\sqrt{n}}$	$\frac{1.22}{\sqrt{n}}$	$\frac{1.36}{\sqrt{n}}$	$\frac{1.37}{\sqrt{n}}$	$\frac{1.63}{\sqrt{n}}$

Gap Test for Random Numbers (cont.)

- The critical value of D in Kolmogorov–Smirnov test is given by

$$D_{0.05} = 1.36 / \sqrt{100} = 0.136$$

Since $D = \max |F(x) - S_N(x)| = 0.0224$ is less than $D_{0.05}$, do not reject the hypothesis of independence on the basis of this test.
(total number of gaps = the number of samples = 100)

Run Test for Random Numbers

➤ Up and Down

- Consider the 40 numbers; both the Kolmogorov-Smirnov and Chi-square would indicate that the numbers are uniformly distributed. But, the sequence does not seem independent.

0.08	0.09	0.23	0.29	0.42	0.55	0.58	0.72	0.89	0.91
0.11	0.16	0.18	0.31	0.41	0.53	0.71	0.73	0.74	0.84
0.02	0.09	0.30	0.32	0.45	0.47	0.69	0.74	0.91	0.95
0.12	0.13	0.29	0.36	0.38	0.54	0.68	0.86	0.88	0.91

—————→
Increasing

Runs Test for Random Numbers (cont.)

- Now, rearrange and there is less reason to doubt independence.*

0.41	0.68	0.89	0.84	0.74	0.91	0.55	0.71	0.36	0.30
0.09	0.72	0.86	0.08	0.54	0.02	0.11	0.29	0.16	0.18
0.88	0.91	0.95	0.69	0.09	0.38	0.23	0.32	0.91	0.53
0.31	0.42	0.73	0.12	0.74	0.45	0.13	0.47	0.58	0.29

Runs Test for Random Numbers (cont.)

- Concerns:

- ✓ Number of runs (direction changes to an increasing or decreasing sequence)
- ✓ Length of runs

- Note.

If N is the number of elements (RNs) in a sequence, the maximum number of runs is $N-1$ (keep changing the direction) and the minimum number of runs is one (w/o changing direction).

Runs Test for Random Numbers (cont.)

- If “a” is the number of runs in a truly random sequence, the mean and variance of “a” is given by

$$\mu_a = (2N - 1) / 3$$

$$\sigma_a^2 = (16N - 29) / 90$$

- For $N > 20$, the distribution of “a” is approximated by a normal distribution, $N(\mu_a, \sigma_a^2)$.
- This approximation can be used to test the independence of numbers from a generator.

$$Z_0 = (a - \mu_a) / \sigma_a : \text{the } Z \text{ score}$$

Runs Test for Random Numbers (cont.)

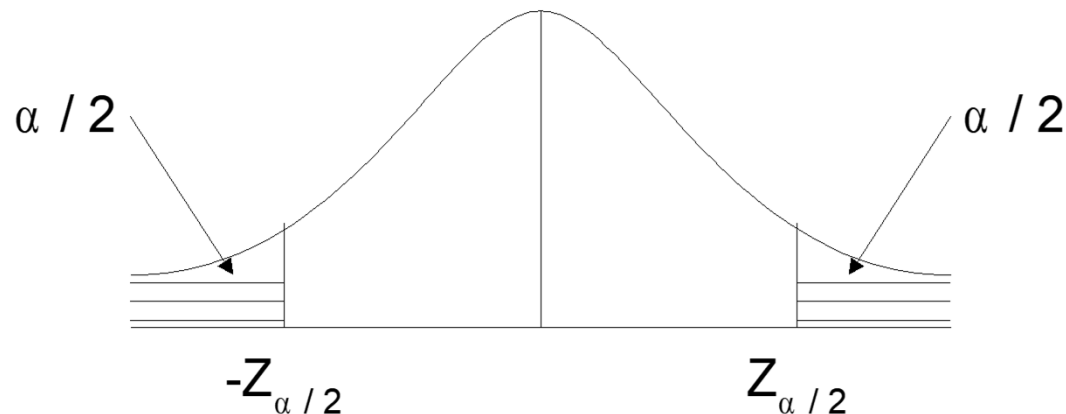
- Substituting for μ_a and $\sigma_a \Rightarrow$

$$Z_a = \{a - [(2N-1)/3]\} / \{\sqrt{(16N-29)/90}\},$$

where $Z \sim N(0,1)$

- Acceptance region for hypothesis of independence

$$-Z_{\alpha/2} \leq Z_0 \leq Z_{\alpha/2}$$



* Too many (“up and then down”s or “down and then up”s) or too small (“monotonously increasing” or “monotonously decreasing”) runs make a reason to doubt!

Runs Test for Random Numbers (cont.)

- Example
 - Based on runs up and runs down, determine whether the following sequence of 40 numbers is such that the hypothesis of independence can be rejected where $\alpha = 0.05$.

0.41	0.68	0.89	0.94	0.74	0.91	0.55	0.62	0.36	0.27
0.19	0.72	0.75	0.08	0.54	0.02	0.01	0.36	0.16	0.28
0.18	0.01	0.95	0.69	0.18	0.47	0.23	0.32	0.82	0.53
0.31	0.42	0.73	0.04	0.83	0.45	0.13	0.57	0.63	0.29

- The sequence of runs up and down is as follows:

(+) + + (-) (+) (-) (+) (-) - - (+) + (-) (+) (-) - (+) (-) (+)
 (-) - (+) (-) - (+) (-) (+) + (-) - (+) + (-) (+) (-) - (+) + (-)

Test for Random Numbers (cont.)

- The sequence of runs up and down is as follows:

+ + + - + - + - - - + + - + - - + - + - - - + - - + - + + - - + + - + - - + + -

There are 26 runs (13 ups + 13 downs) in this sequence.

With $N=40$ and $a=26$,

$$\mu_a = \{2(40) - 1\} / 3 = 26.33 \quad \text{and}$$

$$\sigma_a^2 = \{16(40) - 29\} / 90 = 6.79$$

$$\text{Then, } Z_0 = (26 - 26.33) / \sqrt{6.79} = -0.13 < Z_{0.025}$$

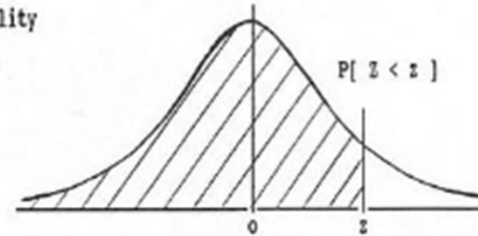
- Now, the critical value is $Z_{0.025} = 1.96$, so the independence of the numbers cannot be rejected on the basis of this test.

STANDARD STATISTICAL TABLES

1. Areas under the Normal Distribution

The table gives the cumulative probability up to the standardised normal value z i.e.

$$P[Z < z] = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}z^2\right) dz$$



| z | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
|-----|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5159 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7854 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8804 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9773 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9865 | 0.9868 | 0.9871 | 0.9874 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9924 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9980 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| z | 3.00 | 3.10 | 3.20 | 3.30 | 3.40 | 3.50 | 3.60 | 3.70 | 3.80 | 3.90 |
| P | 0.9986 | 0.9990 | 0.9993 | 0.9995 | 0.9997 | 0.9998 | 0.9998 | 0.9999 | 0.9999 | 1.0000 |

$$Z_{0.025} = 1.96$$

Review

Normal Distribution

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}, \text{ for } -\infty < x < \infty.$$

✓ Let x_1, \dots, x_n be independent and identically distributed as $M(\mu, \sigma^2)$, i.e. a sample of size n from a normally distributed population with expected value μ and variance σ^2 .

✓ Let $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$ be the sample mean (normally distributed) and

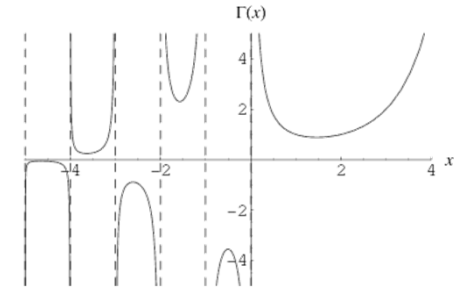
let $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ be the sample variance (χ^2 -distributed).

✓ Then the random variable $\frac{\bar{x} - \mu}{(\sigma/\sqrt{n})}$ follows a standard normal distribution (with expected value 0 and variance 1), and the random variable $\frac{\bar{x} - \mu}{(s/\sqrt{n})}$ follows a Student's t -distribution with $n - 1$ degrees of freedom.

(Student's) t Distribution

penname of the developer, William S. Gosset

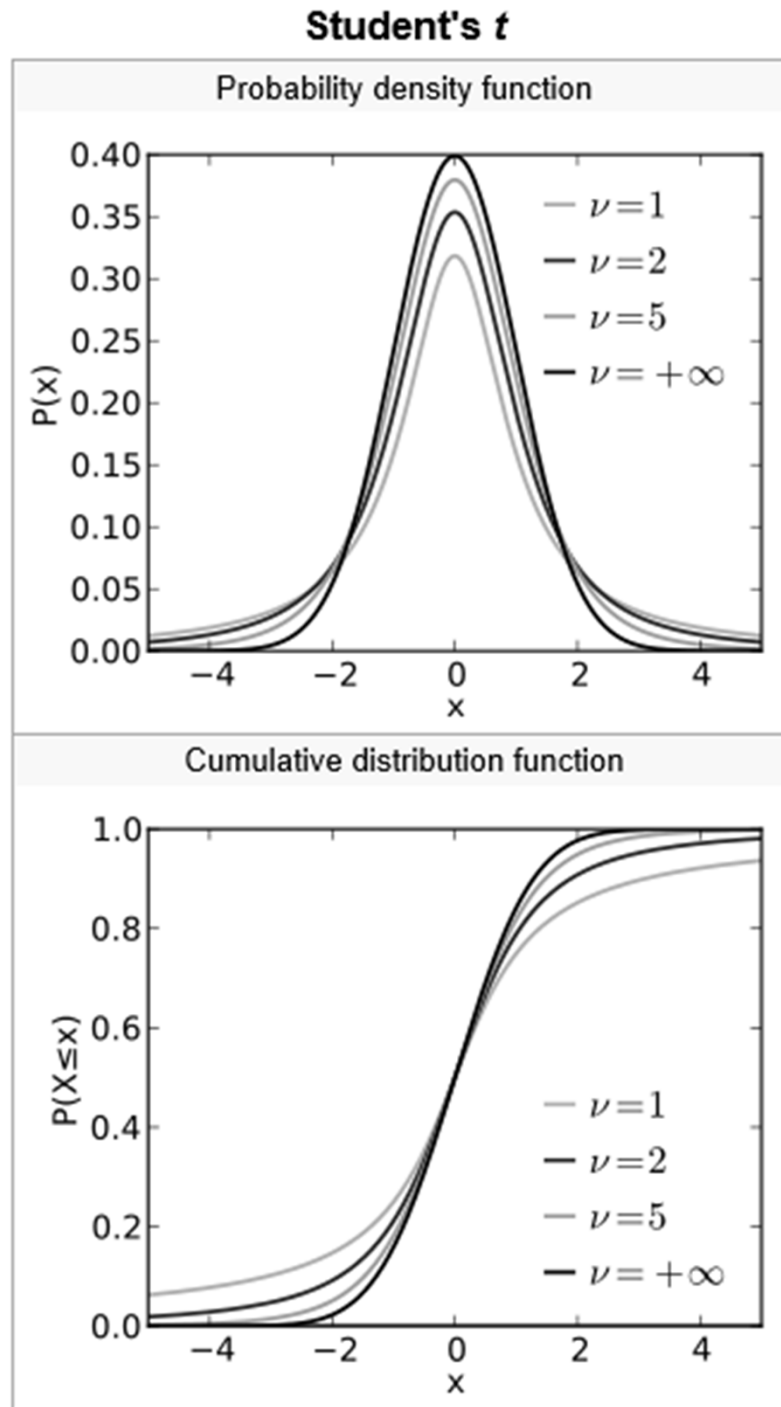
$$f(t) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi} \Gamma\left(\frac{\nu}{2}\right)} \left(1 + \frac{t^2}{\nu}\right)^{-\frac{\nu+1}{2}}$$



where ν = the number of degrees of freedom and Γ is the gamma function. $\Gamma(n) = (n-1)!$ for positive integer n : $\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$

- The function is symmetric, and its overall shape resembles the bell shape of a normally distributed variable with mean 0 and variance 1, except that it is a bit lower and wider.
- As the number of degrees of freedom, ν , grows, the t -distribution approaches the normal distribution with mean 0 and variance 1 (or standard normal distribution). For this reason ν is also known as the normality parameter.

Review



- The normal distribution is shown as a black line ($\nu = \infty$) for comparison.
- T -distribution becomes closer to the normal distribution as ν increases.
- The standard Cauchy distribution is a special case of t -distribution with $\nu = n-1 = 1$:

$$f(t) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi} \Gamma\left(\frac{\nu}{2}\right)} \left(1 + \frac{t^2}{\nu}\right)^{-\frac{\nu+1}{2}}$$

$$\begin{aligned} f(t)\big|_{\nu=1} &= \frac{\Gamma\left(\frac{1+1}{2}\right)}{\sqrt{\pi} \Gamma\left(\frac{1}{2}\right)} (1 + t^2)^{-\frac{1+1}{2}} \\ &= \frac{1}{\pi(1+t^2)} = f(t; 0, 1) \end{aligned}$$

t-distribution
table

| cum. prob | t _{.50} | t _{.75} | t _{.80} | t _{.85} | t _{.90} | t _{.95} | t _{.975} | t _{.99} | t _{.995} | t _{.999} | t _{.9995} |
|-----------|------------------|------------------|------------------|------------------|------------------|------------------|-------------------|------------------|-------------------|-------------------|--------------------|
| one-tail | 0.50 | 0.25 | 0.20 | 0.15 | 0.10 | 0.05 | 0.025 | 0.01 | 0.005 | 0.001 | 0.0005 |
| two-tails | 1.00 | 0.50 | 0.40 | 0.30 | 0.20 | 0.10 | 0.05 | 0.02 | 0.01 | 0.002 | 0.001 |
| df | | | | | | | | | | | |
| 1 | 0.000 | 1.000 | 1.376 | 1.963 | 3.078 | 6.314 | 12.71 | 31.82 | 63.66 | 318.31 | 636.62 |
| 2 | 0.000 | 0.816 | 1.061 | 1.386 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.327 | 31.599 |
| 3 | 0.000 | 0.765 | 0.978 | 1.250 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.215 | 12.924 |
| 4 | 0.000 | 0.741 | 0.941 | 1.190 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 0.000 | 0.727 | 0.920 | 1.156 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 0.000 | 0.718 | 0.906 | 1.134 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 0.000 | 0.711 | 0.896 | 1.119 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 0.000 | 0.706 | 0.889 | 1.108 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 0.000 | 0.703 | 0.883 | 1.100 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 0.000 | 0.700 | 0.879 | 1.093 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 0.000 | 0.697 | 0.876 | 1.088 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 0.000 | 0.695 | 0.873 | 1.083 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 0.000 | 0.694 | 0.870 | 1.079 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 0.000 | 0.692 | 0.868 | 1.076 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 0.000 | 0.691 | 0.866 | 1.074 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 0.000 | 0.690 | 0.865 | 1.071 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 0.000 | 0.689 | 0.863 | 1.069 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 0.000 | 0.688 | 0.862 | 1.067 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 0.000 | 0.688 | 0.861 | 1.066 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 0.000 | 0.687 | 0.860 | 1.064 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 0.000 | 0.686 | 0.859 | 1.063 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 0.000 | 0.686 | 0.858 | 1.061 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 0.000 | 0.685 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.768 |
| 24 | 0.000 | 0.685 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |
| 25 | 0.000 | 0.684 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060 | 2.485 | 2.787 | 3.450 | 3.725 |
| 26 | 0.000 | 0.684 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056 | 2.479 | 2.779 | 3.435 | 3.707 |
| 27 | 0.000 | 0.684 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052 | 2.473 | 2.771 | 3.421 | 3.690 |
| 28 | 0.000 | 0.683 | 0.855 | 1.056 | 1.313 | 1.701 | 2.048 | 2.467 | 2.763 | 3.408 | 3.674 |
| 29 | 0.000 | 0.683 | 0.854 | 1.055 | 1.311 | 1.699 | 2.045 | 2.462 | 2.756 | 3.396 | 3.659 |
| 30 | 0.000 | 0.683 | 0.854 | 1.055 | 1.310 | 1.697 | 2.042 | 2.457 | 2.750 | 3.385 | 3.646 |
| 40 | 0.000 | 0.681 | 0.851 | 1.050 | 1.303 | 1.684 | 2.021 | 2.423 | 2.704 | 3.307 | 3.551 |
| 60 | 0.000 | 0.679 | 0.848 | 1.045 | 1.296 | 1.671 | 2.000 | 2.390 | 2.660 | 3.232 | 3.460 |
| 80 | 0.000 | 0.678 | 0.846 | 1.043 | 1.292 | 1.664 | 1.990 | 2.374 | 2.639 | 3.195 | 3.416 |
| 100 | 0.000 | 0.677 | 0.845 | 1.042 | 1.290 | 1.660 | 1.984 | 2.364 | 2.626 | 3.174 | 3.390 |
| 1000 | 0.000 | 0.675 | 0.842 | 1.037 | 1.282 | 1.646 | 1.962 | 2.330 | 2.581 | 3.098 | 3.300 |
| Z | 0.000 | 0.674 | 0.842 | 1.036 | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 | 3.090 | 3.291 |
| | 0% | 50% | 60% | 70% | 80% | 90% | 95% | 98% | 99% | 99.8% | 99.9% |
| | Confidence Level | | | | | | | | | | |

Poker Test for Random Numbers (cont.)

- *based on the frequency with which certain digits are repeated.*

Example:

0.2(55) 0.5(77) 0.(33)1 0.4(14) 0.8(28) 0.9(09)

Note: a pair of like digits appear in each number generated.

Poker Test for Random Numbers (cont.)

➤ In 3-digit numbers, there are only 3 possibilities.

- $P(3 \text{ different digits})$
 $= (2^{\text{nd}} \text{ diff. from } 1^{\text{st}}) * P(3^{\text{rd}} \text{ diff. from } 1^{\text{st}} \text{ \& } 2^{\text{nd}})$
 $= (0.9) (0.8) = 0.72$
- $P(3 \text{ like digits})$
 $= (2^{\text{nd}} \text{ digit same as } 1^{\text{st}}) * P(3^{\text{rd}} \text{ digit same as } 1^{\text{st}})$
 $= (0.1) (0.1) = 0.01$
- $P(\text{exactly one pair}) = 1 - 0.72 - 0.01 = 0.27$ or
 $= (2^{\text{nd}} \text{ digit same as } 1^{\text{st}}) * (3^{\text{rd}} \text{ diff. from } 1^{\text{st}})$
 $+ (2^{\text{nd}} \text{ digit diff. from } 1^{\text{st}}) * (3^{\text{rd}} \text{ same as } 1^{\text{st}} \text{ or } 2^{\text{nd}})$
 $= (0.1) (0.9) + (0.9) (0.2) = 0.27$

Poker Test for Random Numbers (cont.)

- *Example*

A sequence of 1000 three-digit numbers has been generated and an analysis indicates that 680 have three different digits, 289 contain exactly one pair of like digits, and 31 contain three like digits. Based on the poker test, are these numbers independent?

Let $\alpha = 0.05$.

The test is summarized in the next table.

Poker Test for Random Numbers (cont.)

Example (cont.)

| Combination,
i | Observed
Frequency,
O_i | Expected
Frequency,
E_i | $\frac{(O_i - E_i)^2}{E_i}$ |
|------------------------|---------------------------------|---------------------------------|-----------------------------|
| Three different digits | 680 | 720 | 2.24 |
| Three like digits | 31 | 10 | 44.10 |
| Exactly one pair | 289 | 270 | 1.33 |
| | <u>1000</u> | <u>1000</u> | <u>47.65</u> |

- The appropriate degrees of freedom are one less than the number of class intervals.*
- Since $\chi^2_{0.05, 2} (= 5.99) < \chi^2 (=47.65)$, the independence of the numbers is rejected on the basis of this test.*

✓ Chi-Squared Distribution (cont.)

chi-square distribution table

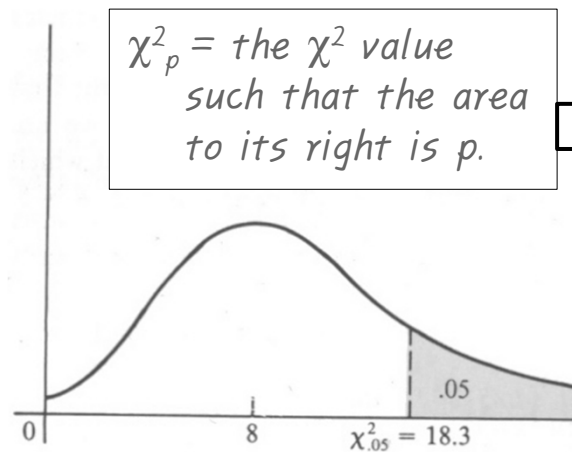


FIGURE 4.5

$P[X_{10}^2 \geq \chi_{.05}^2] = .05$ and $P[X_{10}^2 < \chi_{.05}^2] = .95$.

| df \ p | 0.995 | 0.975 | 0.9 | 0.5 | 0.1 | 0.05 | 0.025 | 0.01 | 0.005 | df |
|--------|-------|-------|-------|--------|--------|--------|--------|--------|--------|----|
| 1 | .000 | .000 | 0.016 | 0.455 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 | 1 |
| 2 | 0.010 | 0.051 | 0.211 | 1.386 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 | 2 |
| 3 | 0.072 | 0.216 | 0.584 | 2.366 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 | 3 |
| 4 | 0.207 | 0.484 | 1.064 | 3.357 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 | 4 |
| 5 | 0.412 | 0.831 | 1.610 | 4.351 | 9.236 | 11.070 | 12.832 | 15.086 | 16.750 | 5 |
| 6 | 0.676 | 1.237 | 2.204 | 5.348 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 | 6 |
| 7 | 0.989 | 1.690 | 2.833 | 6.346 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 | 7 |
| 8 | 1.344 | 2.180 | 3.490 | 7.344 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 | 8 |
| 9 | 1.735 | 2.700 | 4.168 | 8.343 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 | 9 |
| 10 | 2.156 | 3.247 | 4.865 | 9.342 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 | 10 |
| 11 | 2.603 | 3.816 | 5.578 | 10.341 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 | 11 |
| 12 | 3.074 | 4.404 | 6.304 | 11.340 | 18.549 | 21.026 | 23.337 | 26.217 | 28.300 | 12 |
| 13 | 3.565 | 5.009 | 7.042 | 12.340 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 | 13 |
| 14 | 4.075 | 5.629 | 7.790 | 13.339 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 | 14 |
| 15 | 4.601 | 6.262 | 8.547 | 14.339 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 | 15 |

- p = the probability that a random sample from a true Poisson distribution would have a larger value of χ^2 than the specified value shown in the table.
 - Very low value (say less than 0.02) indicate abnormally large fluctuations in the data whereas very high probabilities (greater than 0.98) indicate abnormally small fluctuation.
 - Perfect fit to the Poisson distribution for large samples would yield a probability 0.50.

Correlation Test for Random Numbers

- ✓ Generate U_1, \dots, U_n
- ✓ Compute an estimate for the (serial correlation)

$$\hat{\rho}_1 = \frac{\sum_{i=1}^{n-1} (U_i - \bar{U}(n))(U_{i+1} - \bar{U}(n))}{\sum_{i=1}^n (U_i - \bar{U}(n))^2}$$

where $\bar{U}(n)$ is the sample mean.

- ✓ If the U_i s are really i.i.d. $U(0,1)$, then $\hat{\rho}_1$ should be close to zero.
- * “Consequential up, down and then up”s makes $\hat{\rho}_1$ be negatively large. Also, “consequential ups” or “consequential downs” make $\hat{\rho}_1$ be positively large!
- ✓ For large n , $p(-2/\sqrt{n} \leq \hat{\rho}_1 \leq 2/\sqrt{n}) \approx 0.95$.

Thus reject H_0 at the 5% confidence level

if $\hat{\rho}_1 < -2/\sqrt{n}$ or $\hat{\rho}_1 > 2/\sqrt{n}$.

Correlation Test for Random Numbers (cont.)

- ✓ Generate u_1, \dots, u_n for the variable U
- ✓ Compute an estimate for the serial (auto-)correlation

$$\hat{\rho}_k = \frac{\sum_{i=1}^{n-k} \{u_i - \bar{U}(n)\} \cdot \{u_{i+k} - \bar{U}(n)\}}{\sum_{i=1}^n \{u_i - \bar{U}(n)\}^2} \quad \text{at lag } k (\geq 1),$$

where $\bar{U}(n)$ is the sample mean.

$$\begin{aligned} \hat{\rho}_k &= \frac{(n-k)}{(n-1)} \cdot \frac{\frac{1}{(n-k)} \sum_{i=1}^{n-k} \{u_i - \bar{U}(n)\} \cdot \{u_{i+k} - \bar{U}(n)\}}{\frac{1}{n-1} \sum_{i=1}^n \{u_i - \bar{U}(n)\}^2} \\ &= \frac{(n-k)}{(n-1)} \cdot \frac{E[\{u_i - \bar{U}(n)\} \cdot \{u_{i+k} - \bar{U}(n)\}]}{s(U)^2} \end{aligned}$$

Correlation Test for Random Numbers (cont.)

$$\begin{aligned}\hat{\rho}_k &= \frac{(n-k)}{(n-1)} \cdot \frac{E[\{u_i - \bar{U}(n)\} \cdot \{u_{i+k} - \bar{U}(n)\}]}{\sigma(U)^2} \\ &= \frac{(n-k)}{(n-1)} \cdot \frac{\text{cov}(U_i, U_{i+k}) + E[\{u_i - \bar{U}(n)\}] \cdot E[\{u_{i+k} - \bar{U}(n)\}]}{\sigma(U)^2}\end{aligned}$$

If $\{U_i\}$ and $\{U_{i+k}\}$ are independent, $\text{cov}(U_i, U_{i+k}) = 0$.

$$\hat{\rho}_k = \frac{(n-k)}{(n-1)} \cdot \frac{\{E[u_i] - \bar{U}(n)\} \cdot \{E[u_{i+k}] - \bar{U}(n)\}}{\sigma(U)^2}$$

$$\hat{\rho}_k \cdot (n-1) = \frac{\{E[u_i] - \bar{U}(n)\}}{\{\sigma(U)/\sqrt{n-k}\}} \cdot \frac{\{E[u_{i+k}] - \bar{U}(n)\}}{\{\sigma(U)/\sqrt{n-k}\}} \sim (\text{Z score}) \cdot (\text{Z score})$$

$\therefore -1.96 \leq (\text{Z score}) \leq 1.96$ occurs with 95% of confidence, or

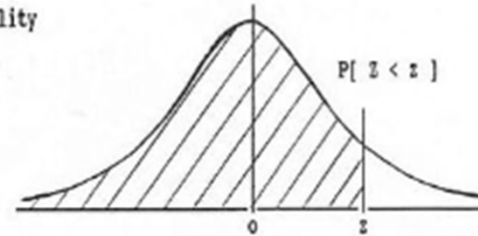
$$-1.96 \leq \sqrt{\hat{\rho}_k \cdot (n-1)} \leq 1.96 \text{ occurs with 95\% of confidence}$$

STANDARD STATISTICAL TABLES

1. Areas under the Normal Distribution

The table gives the cumulative probability up to the standardised normal value z i.e.

$$P[Z < z] = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}z^2\right) dz$$



| z | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
|-----|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5159 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7854 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8804 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9773 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9865 | 0.9868 | 0.9871 | 0.9874 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9924 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9980 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| z | 3.00 | 3.10 | 3.20 | 3.30 | 3.40 | 3.50 | 3.60 | 3.70 | 3.80 | 3.90 |
| P | 0.9986 | 0.9990 | 0.9993 | 0.9995 | 0.9997 | 0.9998 | 0.9998 | 0.9999 | 0.9999 | 1.0000 |

$$Z_{0.025} = 1.96$$

Review

Covariance and Correlation Coefficient

- The covariance is defined by

$$\text{covariance} = \text{cov}(x, y) = \overline{xy} - \bar{x} \cdot \bar{y} \quad (20)$$

- If x and y are independent, $\text{cov}(x, y) = 0$.

sufficient cond. necessary cond.

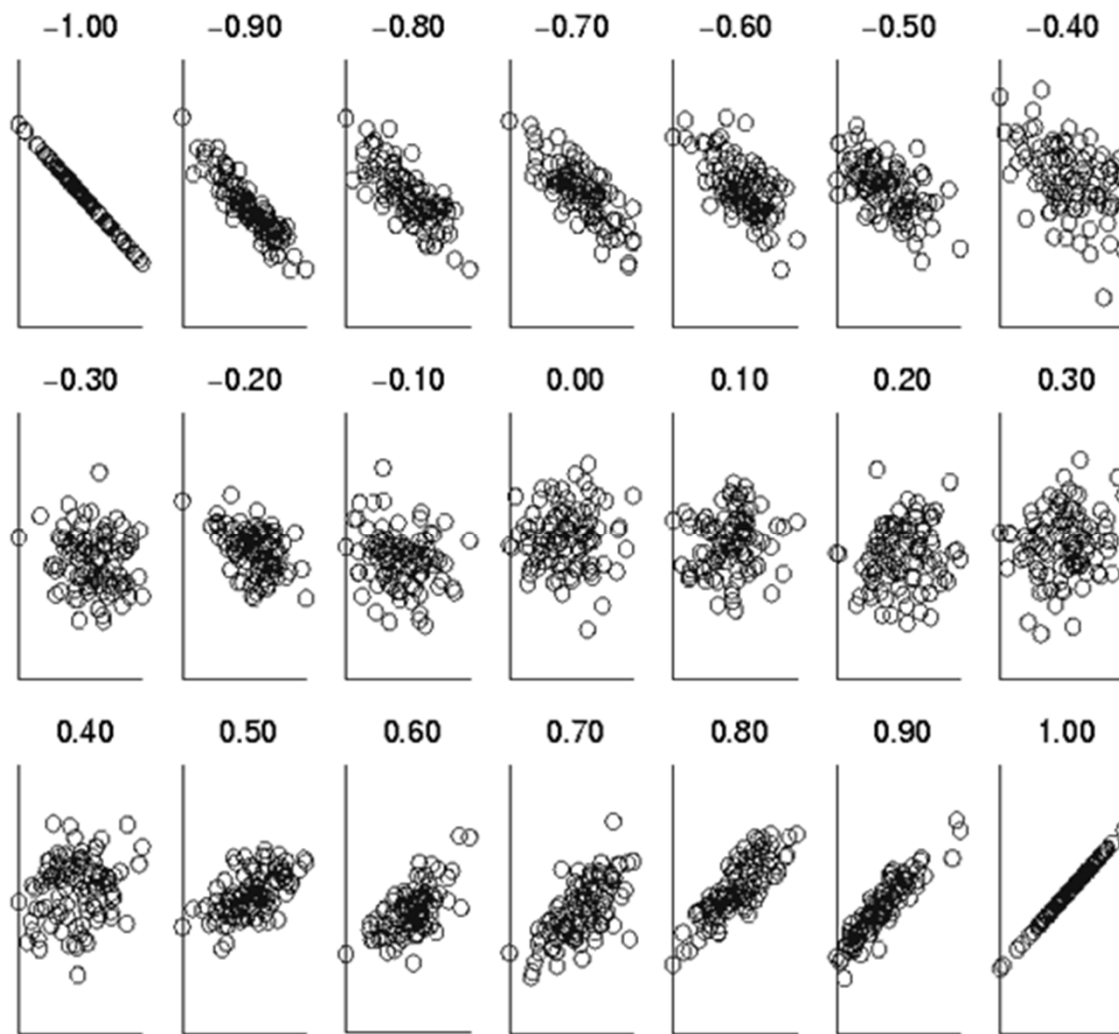
- It is possible to have $\text{cov}(x, y) = 0$ even if x and y are not independent.
- The covariance can be negative.

- Correlation coefficient is a convenient measure of the degree to which two random variables are correlated (or anti-correlated).

$$\text{correlation coefficient} = \rho(x, y) \equiv \text{cov}(x, y) / [\sigma^2(x) \cdot \sigma^2(y)]^{1/2} \quad (21)$$

where $-1 \leq \rho(x, y) \leq 1$.

Review



$\rho = 1$ when $y = ax + b$
 $\rho = -1$ when $y = -ax + b$
($a > 0$)

$\rho = \text{correlation coefficient}$