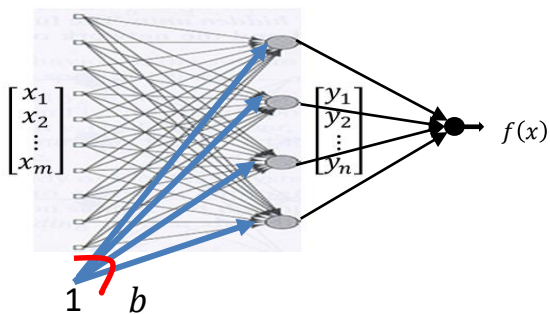


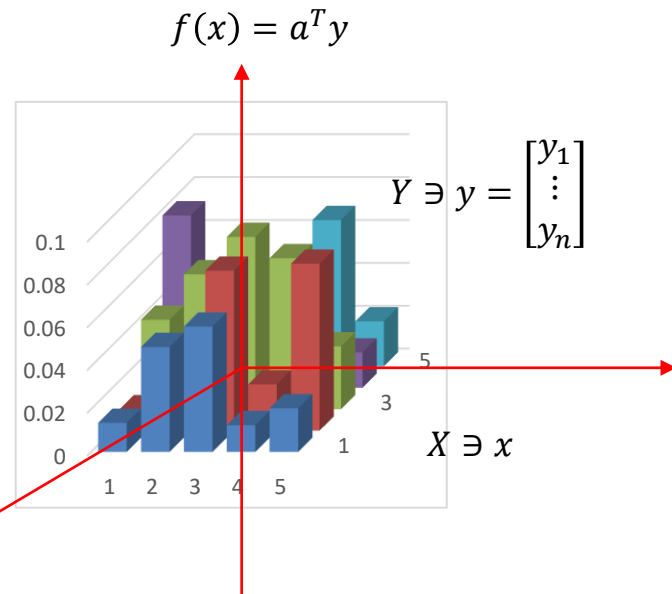
Nonlinear Mapping

$$T: X \rightarrow Y, \quad y = \sigma(Wx + b), \quad f(x) = a^T y$$



Universal Approximation Theorem Let ξ be a non-constant, bounded, and monotonically-increasing continuous activation function, $f: [0, 1]^d \rightarrow \mathbb{R}$ continuous function, and $\epsilon > 0$. Then, $\exists n$ and parameters $a, b \in \mathbb{R}^n$, $W \in \mathbb{R}^{n \times d}$ s.t.

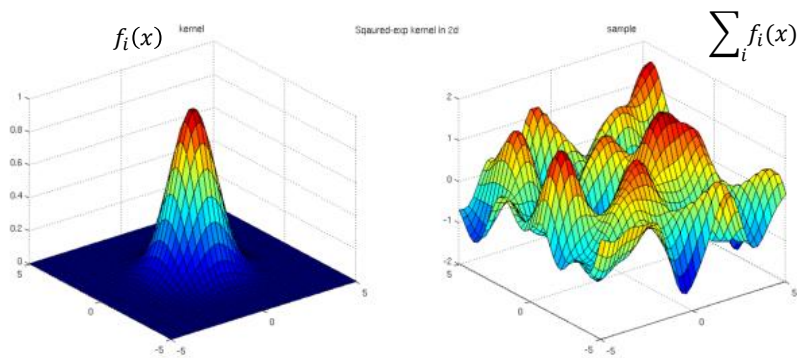
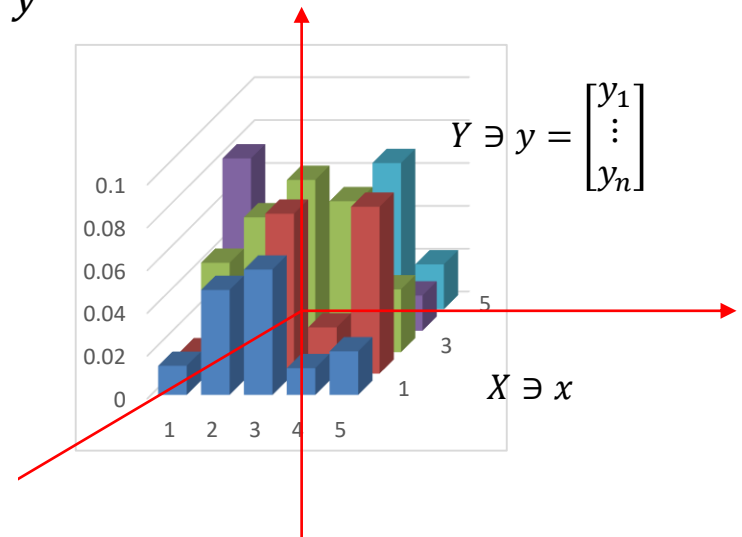
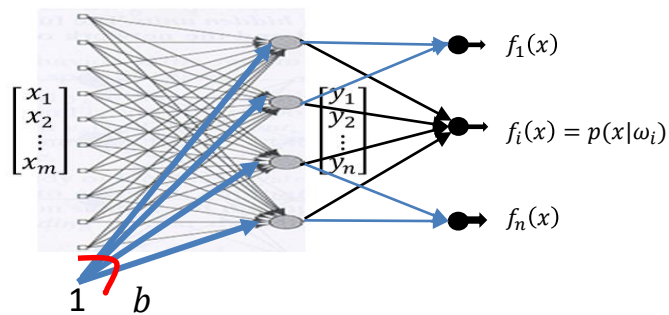
$$\left| \sum_{i=1}^n a_i \xi(w_i^T x + b_i) - f(x) \right| < \epsilon \quad \forall x \in [0, 1]^d$$



[Geometric Deep Learning on graph and manifolds](#), Michael Bronstein, SIAM 2018, Imperial College London

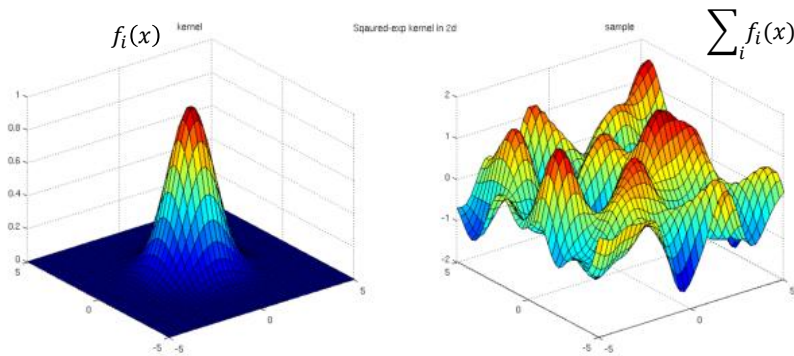
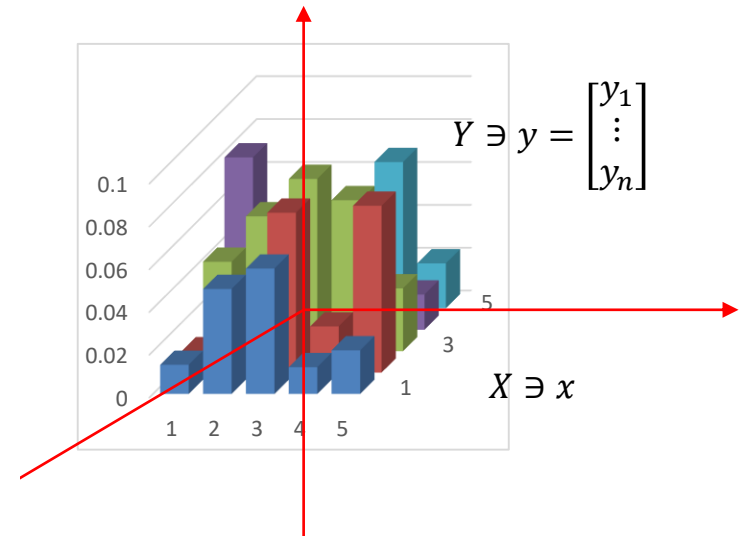
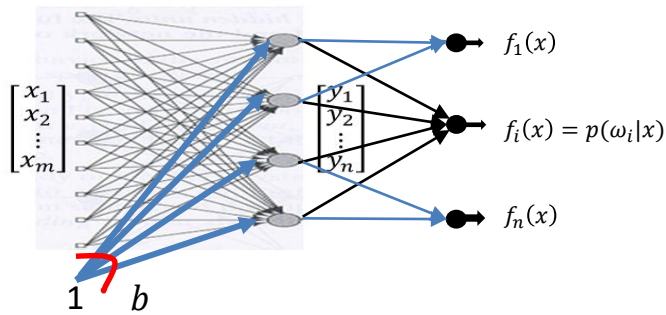
Nonlinear Mapping

$$T: X \rightarrow Y, \quad y = \sigma(Wx + b), \quad f_i(x) = a_i^T y$$



Nonlinear Mapping

$$T: X \rightarrow Y, y = \sigma(Wx + b), f_i(x) = \frac{a_i^T y}{\sum_j a_j^T y} \text{ (softmax)}$$



Feature Dimension Reduction: PCA & LDA (I)

Jin Young Choi

Seoul National University

Outline

Feature Extraction

Introduction of PCA & LDA

Principal Component Analysis (PCA)

Linear Discriminant Analysis (FLDA)

Multiple Discriminant Analysis (MDA)

Simple Enhancement of PCA/LDA

Feature Extraction

- Features

 - Weight, Height, Width, Volume, Head size, ...

 - Edge, Shape, Geometric Relations ...

 - RGB Color for each pixel

 - SIFT, SURF, HOG, ...

- Feature Extraction from Raw Data

 - Pixel Valued Vector is raw data vector

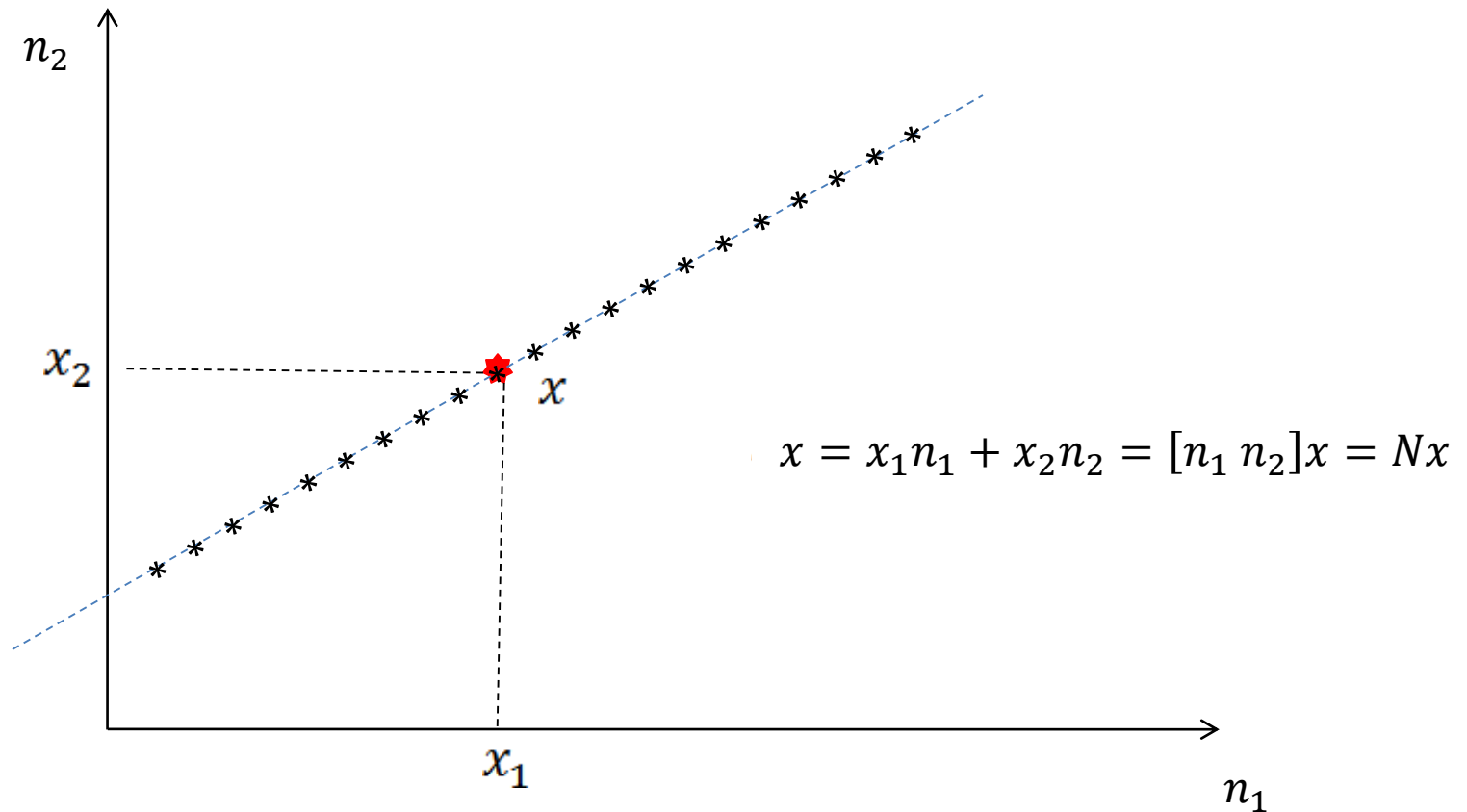
 - Raw data vector is redundant

 - The dimension should be reduced

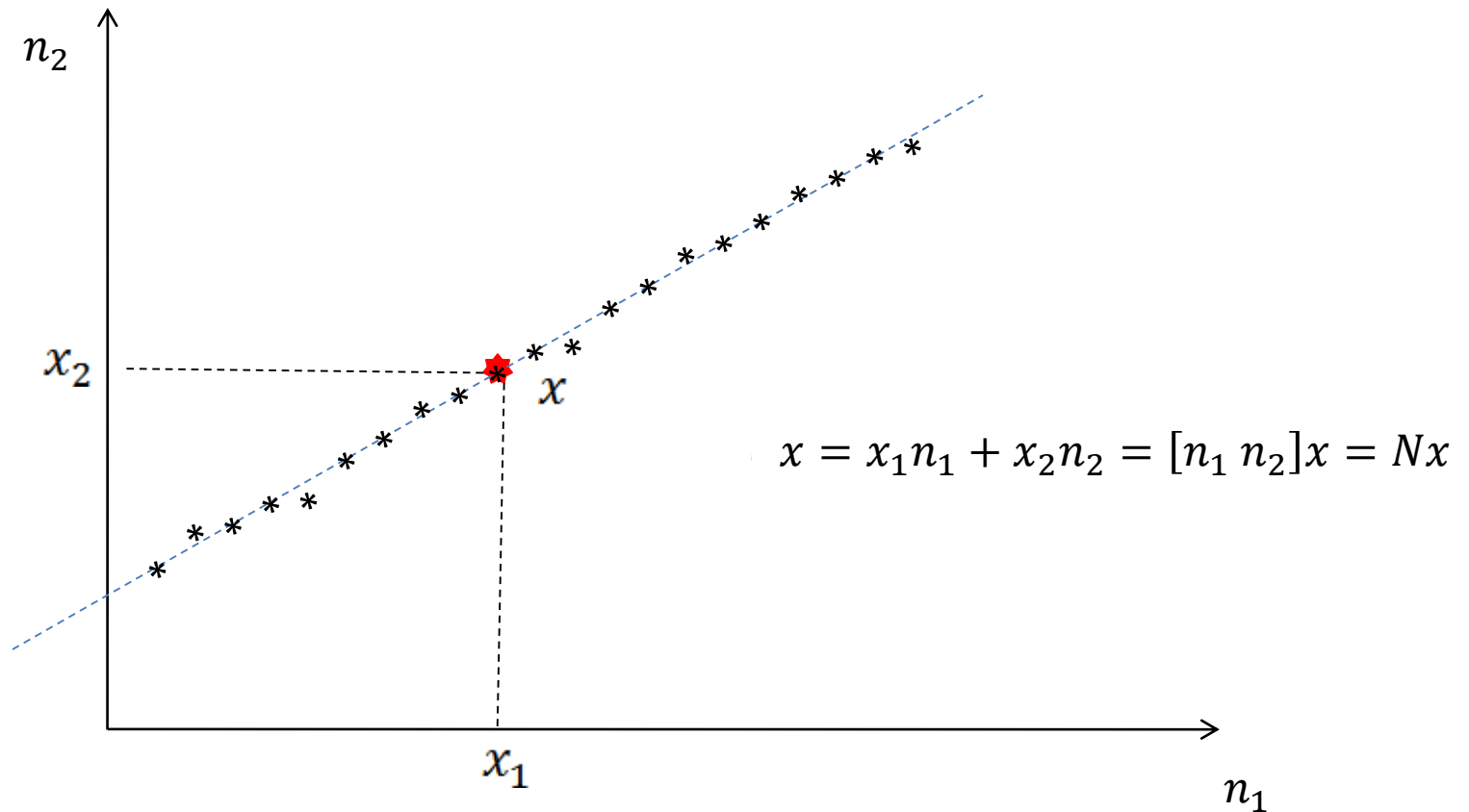
Component Analysis and Discriminants

- How to reduce excessive dimensionality?
 - Answer: Combine features highly dependent to each other.
 - Linear methods project high-dimensional data onto lower dimensional space.
 - Principal Components Analysis (PCA)
 - seeks the projection which best represents the data in a least-square error sense.
 - Linear Discriminant Analysis (LDA) or Fisher Linear Discriminant
 - seeks the projection that best separates the data in a least-square discrimination error sense.
-

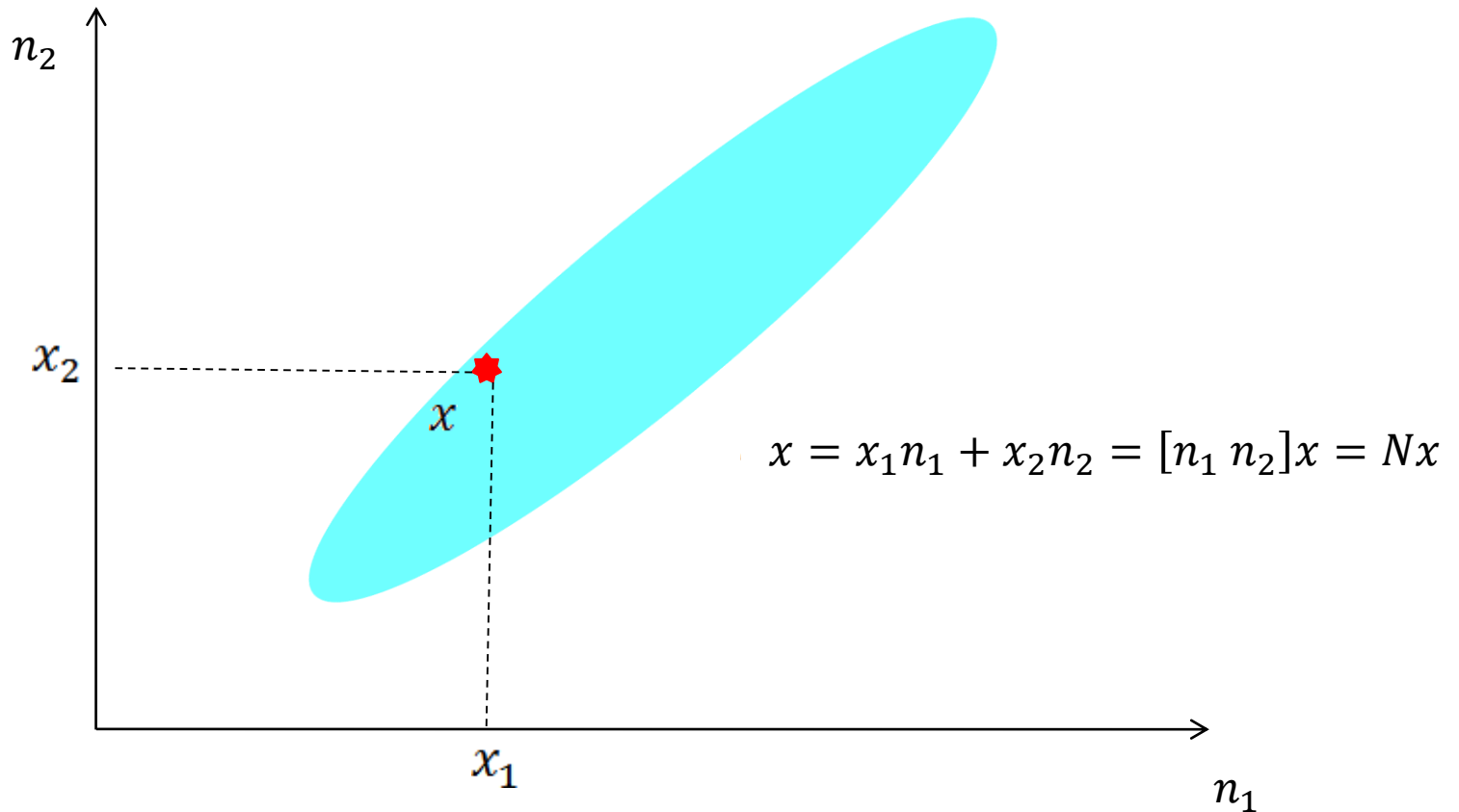
Principal Component Analysis



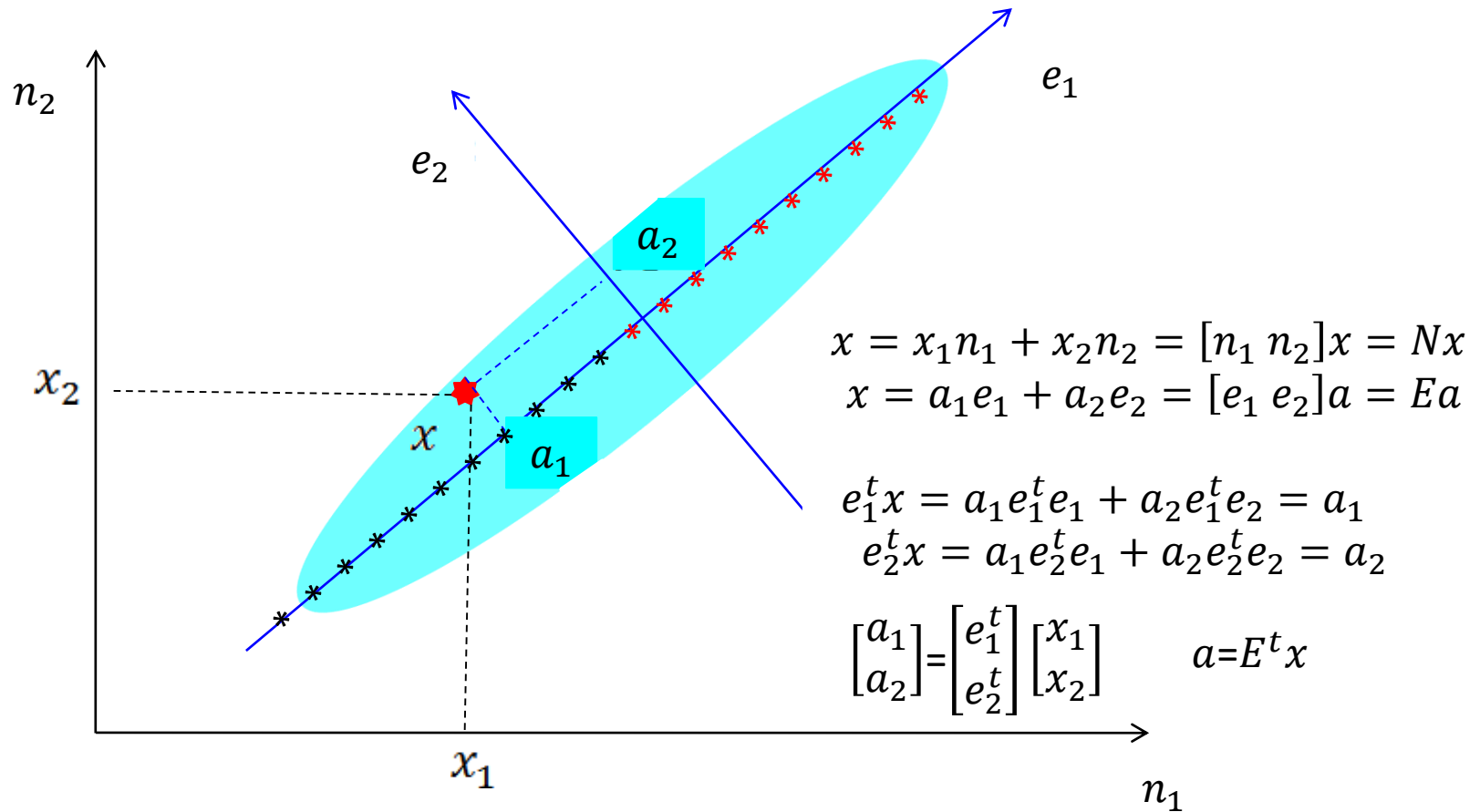
Principal Component Analysis



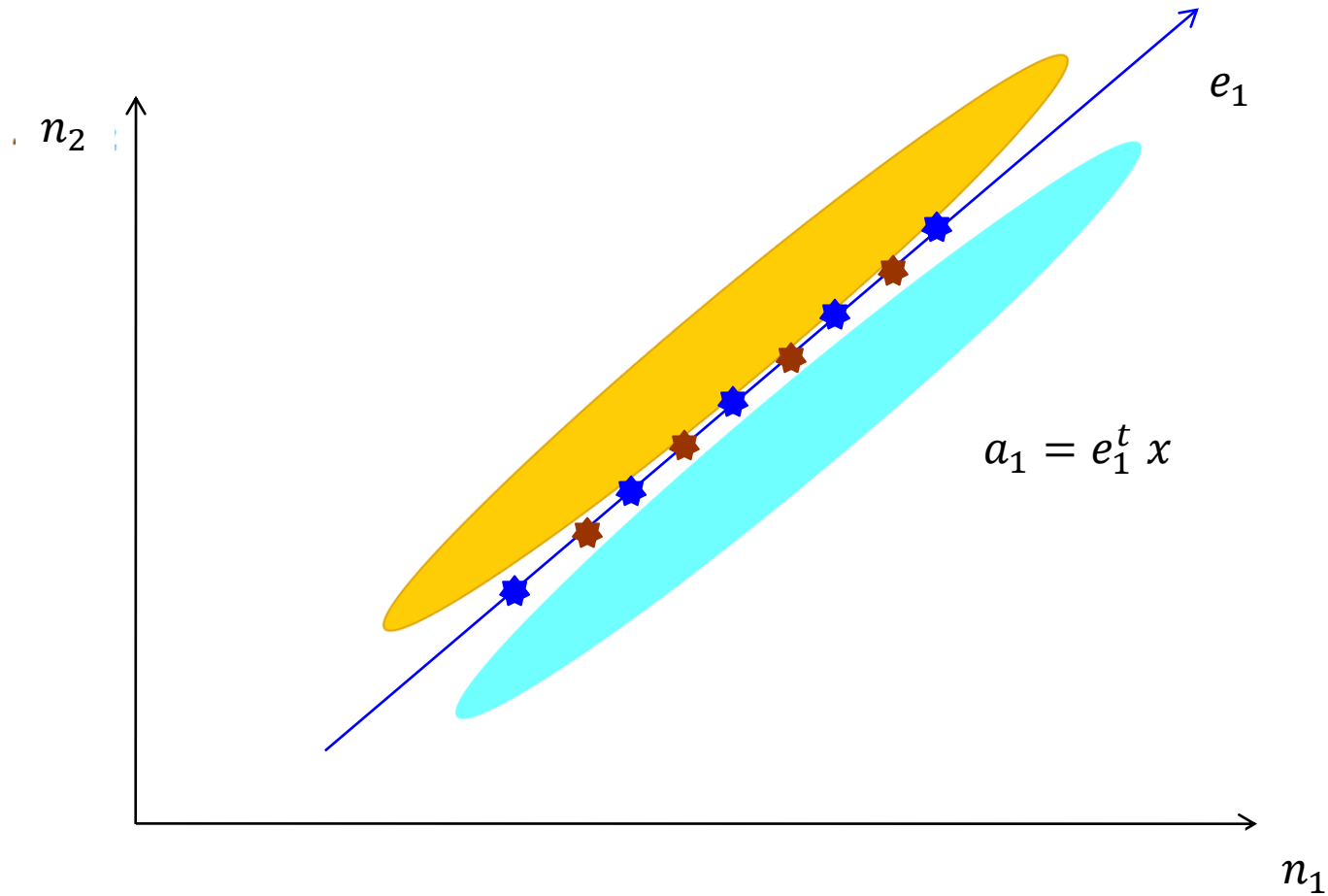
Principal Component Analysis



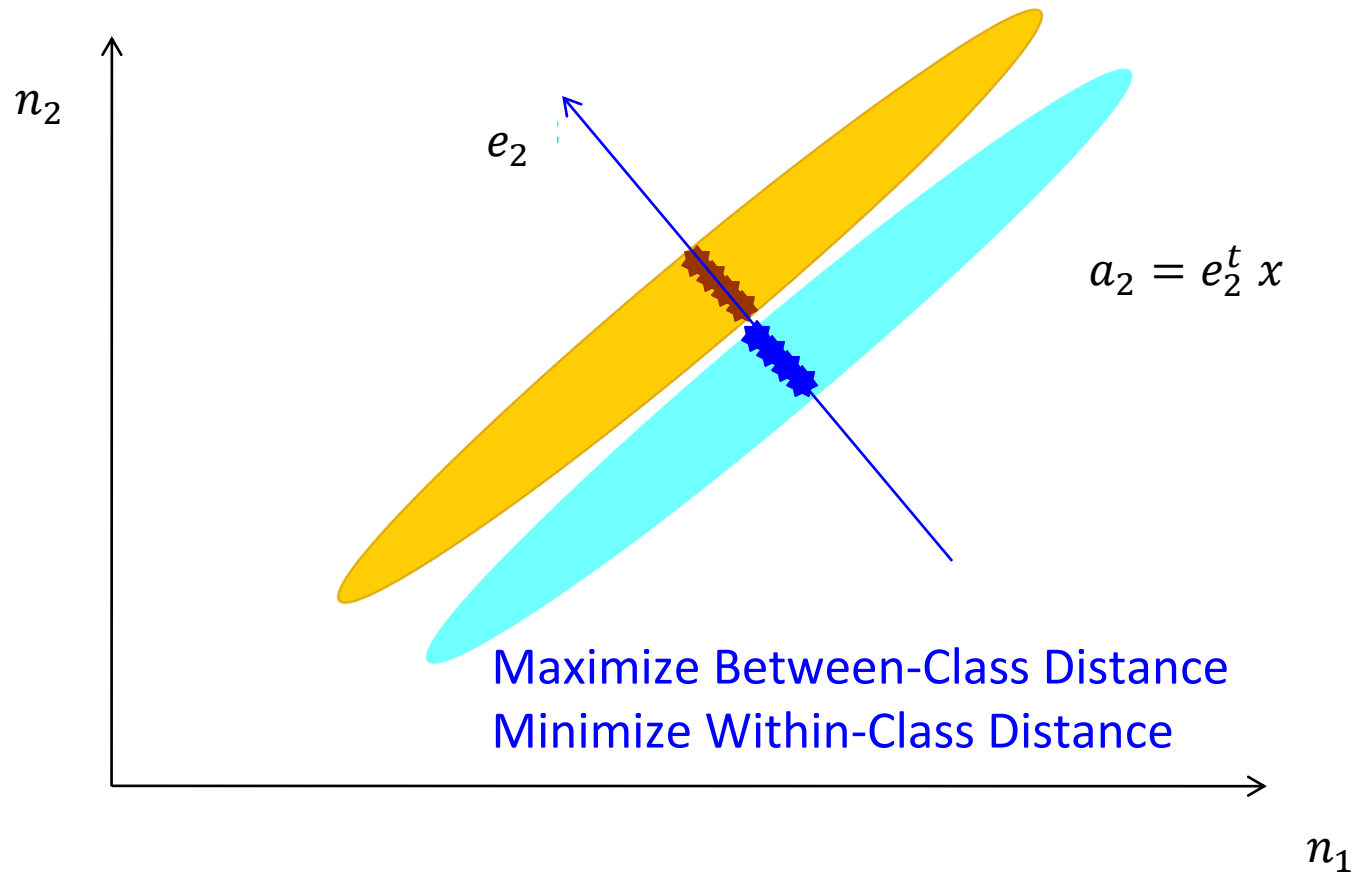
Principal Component Analysis



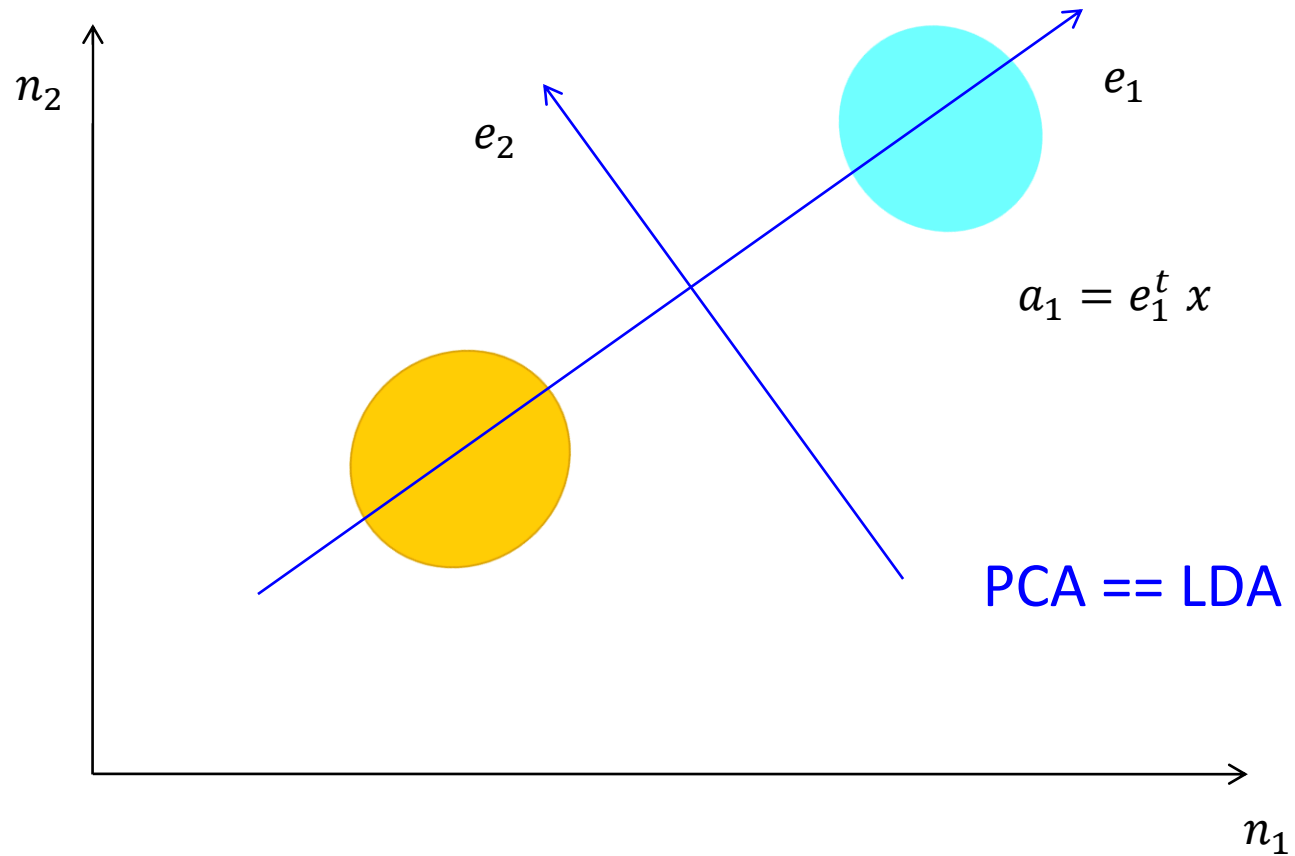
Linear Discriminant Analysis



Linear Discriminant Analysis



PCA & LDA



Principal Components Analysis (PCA)

- How to represent n d -dimensional vector samples $\{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ by a single vector \mathbf{x}_0 ?
 - Find \mathbf{x}_0 that minimizes squared error correction function

$$J_0(\mathbf{x}_0) = \sum_{k=1}^n \|\mathbf{x}_0 - \mathbf{x}_k\|^2 .$$

Principal Components Analysis (PCA)

- How to represent n d -dimensional vector samples $\{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ by a single vector \mathbf{x}_0 ?

- Find \mathbf{x}_0 that minimizes squared error correction function

$$J_0(\mathbf{x}_0) = \sum_{k=1}^n \|\mathbf{x}_0 - \mathbf{x}_k\|^2 .$$

- The solution is sample mean

$$\mathbf{x}_0 = \mathbf{m} = \frac{1}{n} \sum_{k=1}^n \mathbf{x}_k$$

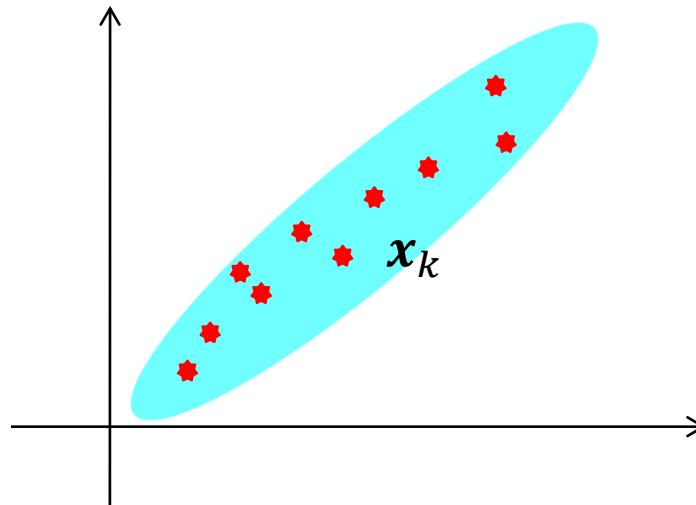
- This is **zero-dimensional** representation of the data set.
 - **One-dimensional** representation by projecting the data onto a line through the sample mean reveals variability in the data.
-

Principal Components Analysis (PCA)

- This is **zero-dimensional** representation of the data set.

$$\mathbf{x}_0 = \mathbf{m} = \frac{1}{n} \sum_{k=1}^n \mathbf{x}_k$$

- **One-dimensional** representation by projecting the data onto a line through the sample mean reveals variability in the data.



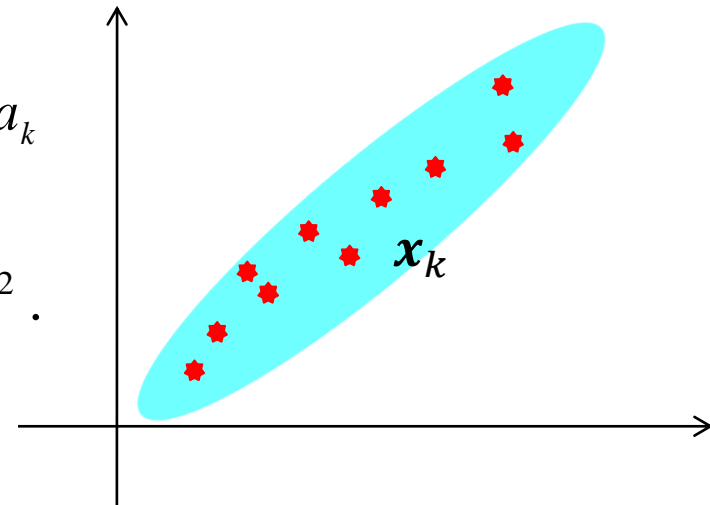
PCA ; Projection

- Let \mathbf{e} be a unit vector in a direction of the line. The equation of the line

$$\mathbf{x} = \mathbf{m} + a \mathbf{e}$$

- Representing \mathbf{x}_k by $\mathbf{m} + a_k \mathbf{e}$ find “optimal” a_k
set minimizing criterion function :

$$J_1(a_1, \dots, a_n, \mathbf{e}) = \sum_{k=1}^n \|\mathbf{m} + a_k \mathbf{e} - \mathbf{x}_k\|^2 .$$



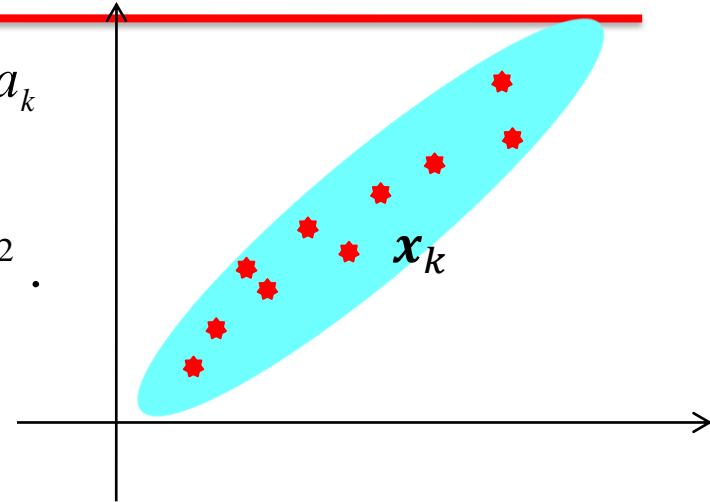
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from $\partial J_1 / \partial a_k = 0$

we find $a_k = \mathbf{e}^t (\mathbf{x}_k - \mathbf{m})$

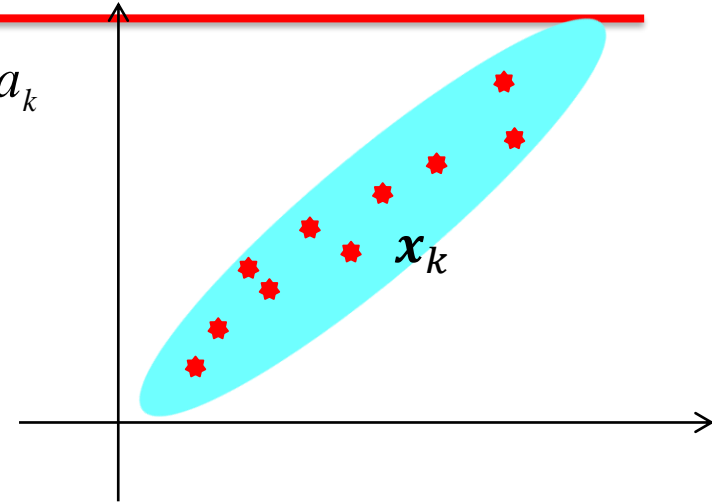


PCA ; Projection

- Representing \mathbf{x}_k by $\mathbf{m} + a_k \mathbf{e}$ find “optimal” a_k

$$a_k = \mathbf{e}^t (\mathbf{x}_k - \mathbf{m})$$

- How to find the *best* direction for \mathbf{e} ?



- **The least square solution:** project the vector \mathbf{x}_k onto the line in the direction of \mathbf{e} , passing through the sample mean.

$$J_1(a_1, \dots, a_n, \mathbf{e}) = \sum_{k=1}^n \|\mathbf{m} + a_k \mathbf{e} - \mathbf{x}_k\|^2 . \quad a_k = \mathbf{e}^t (\mathbf{x}_k - \mathbf{m})$$

- Minimize J w.r.t \mathbf{e} .
-

PCA ; Scatter matrix

- Substituting a_k into $J_1(a, \mathbf{e})$ we find

$$\begin{aligned} J_1(a, \mathbf{e}) &= \sum_{k=1}^n a_k^2 \|\mathbf{e}\|^2 - 2 \sum_{k=1}^n a_k \mathbf{e}^t (\mathbf{x}_k - \mathbf{m}) + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2 \\ &= \sum_{k=1}^n a_k^2 - 2 \sum_{k=1}^n a_k^2 + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2 = - \sum_{k=1}^n [\mathbf{e}^t (\mathbf{x}_k - \mathbf{m})]^2 + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2 \\ &= - \sum_{k=1}^n \mathbf{e}^t (\mathbf{x}_k - \mathbf{m}) (\mathbf{x}_k - \mathbf{m})^t \mathbf{e} + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2 \\ &= -\mathbf{e}^t \mathbf{S} \mathbf{e} + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2 \end{aligned}$$

- where a *scatter matrix* \mathbf{S} which is $(n - 1)$ times of sample covariance matrix

$$\mathbf{S} = \sum_{k=1}^n (\mathbf{x}_k - \mathbf{m})(\mathbf{x}_k - \mathbf{m})^t.$$

PCA ; Scatter matrix

$$J_1(a, \mathbf{e}) = -\mathbf{e}^t \mathbf{S} \mathbf{e} + \sum_{k=1}^n \|\mathbf{x}_k - \mathbf{m}\|^2$$

- Vector \mathbf{e} that minimizes J_1 also maximizes $\mathbf{e}^t \mathbf{S} \mathbf{e}$.
- So we find \mathbf{e} , which maximize $\mathbf{e}^t \mathbf{S} \mathbf{e}$

subject to constraint $\|\mathbf{e}\|=1$

- Let λ be Lagrange multiplier. $L = \mathbf{e}^t \mathbf{S} \mathbf{e} - \lambda(\mathbf{e}^t \mathbf{e} - 1)$
- Differentiating L with respect to \mathbf{e} : $\partial L / \partial \mathbf{e} = 2\mathbf{S} \mathbf{e} - 2\lambda \mathbf{e}$
- By setting to zero we see that \mathbf{e} is an eigenvector of \mathbf{S} :

$$\mathbf{S} \mathbf{e} = \lambda \mathbf{e} \quad \mathbf{e}^t \mathbf{S} \mathbf{e} = \lambda$$

- So to maximize $\mathbf{e}^t \mathbf{S} \mathbf{e}$ takes maximal λ
-

PCA ; Scatter matrix

- The result is easily extended to d' dimensional projection:

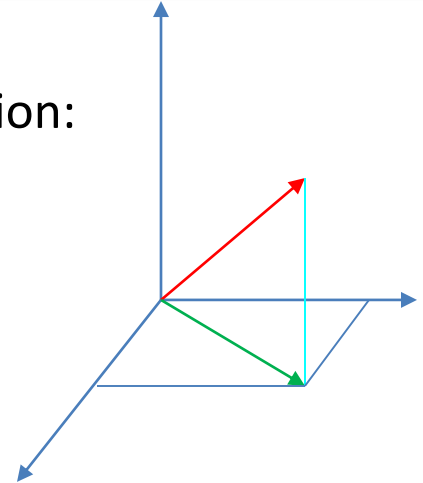
$$\mathbf{x}'_k = \mathbf{m} + \sum_{i=1}^{d'} a_k^i \mathbf{e}_i \quad \text{where} \quad d' \leq d$$

- The criterion function

$$J_{d'} = \sum_{k=1}^n \left\| \left(\mathbf{m} + \sum_{i=1}^{d'} a_k^i \mathbf{e}_i \right) - \mathbf{x}_k \right\|^2$$

is minimized when vectors $\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_{d'}$ are the eigenvectors having the largest eigenvalues.

- The coefficients $a_k^i = \mathbf{e}_i^t (\mathbf{x}_k - \mathbf{m})$ are *principal components*.



Error function

- If $d' < d$ error which is made by dropping the last terms is

$$\begin{aligned} J_{d'} &= \sum_{k=1}^n \left\| \sum_{i=d'+1}^d a_k^i \mathbf{e}_i \right\|^2 \\ &= \sum_{i=d'+1}^d \mathbf{e}_i^t \sum_{k=1}^n (\mathbf{x}_k - \mathbf{m})(\mathbf{x}_k - \mathbf{m})^t \mathbf{e}_i \\ &= \sum_{i=d'+1}^d \mathbf{e}_i^t \mathbf{S} \mathbf{e}_i = \sum_{i=d'+1}^d \lambda_i \end{aligned}$$

$$\mathbf{x}'_k = \mathbf{m}_k + \sum_{i=1}^{d'} a_k^i \mathbf{e}_i$$

$$a_k^i = \mathbf{e}_i^t (\mathbf{x}_k - \mathbf{m})$$

- This is a sum of lowest eigenvalues.
-

PCA – the algorithm

- Input: $X^{(n)} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$, $\mathbf{x}_k = \langle x_1^k, \dots, x_d^k \rangle$
- Take $d' < d$
- Output: $A^{(n)} = \{\mathbf{a}_1, \dots, \mathbf{a}_n\}$ $\mathbf{a}_k = \{a_1^k, \dots, a_{d'}^k\}$
- Algorithm:
 - Compute the mean of the training set $\mathbf{m} = \frac{1}{n} \sum_{k=1}^n \mathbf{x}_k$.
 - Compute the scatter matrix \mathbf{S} .
 - Find eigenvectors of \mathbf{S} and corresponding eigenvalues:

$$S\{\mathbf{e}_i, \lambda_i\}_{i=1}^d, \quad \forall i: \mathbf{S}\mathbf{e}_i = \lambda_i \mathbf{e}_i, \quad \lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$$

- Choose d' eigenvectors, and for each sample \mathbf{x}_k point compute

$$\mathbf{a}_k = \{\mathbf{e}_i^t (\mathbf{x}_k - \mathbf{m})\}_{i=1}^{d'}$$

Interim Summary

- Principal Component Analysis

- ✓ Feature Extraction
- ✓ Dimension Reduction

$$J_{d'} = \sum_{k=1}^n \left\| \left(\mathbf{m} + \sum_{i=1}^{d'} a_k^i \mathbf{e}_i \right) - \mathbf{x}_k \right\|^2$$

$$\mathbf{S} = \sum_{k=1}^n (\mathbf{x}_k - \mathbf{m})(\mathbf{x}_k - \mathbf{m})^t.$$

$$\mathbf{S}\mathbf{e} = \lambda \mathbf{e}$$

$$\mathbf{e}^t \mathbf{S} \mathbf{e} = \lambda$$

$$a_k^i = \mathbf{e}_i^t (\mathbf{x}_k - \mathbf{m}), i = 1, \dots, d'$$

$$\begin{bmatrix} a_k^1 \\ a_k^2 \\ \dots \\ a_k^{d'} \end{bmatrix} = \begin{bmatrix} \mathbf{e}_1^t \\ \mathbf{e}_2^t \\ \dots \\ \mathbf{e}_{d'}^t \end{bmatrix} (\mathbf{x}_k - \mathbf{m})$$

$$\mathbf{a}_k = \mathbf{E}^t (\mathbf{x}_k - \mathbf{m})$$

$$cf) \quad y_k = \mathbf{W}^t (\mathbf{x}_k - \mathbf{m})$$

Feature Dimension Reduction: PCA & LDA (II)

Jin Young Choi

Seoul National University

Outline

Feature Extraction

Introduction of PCA & LDA

Principal Component Analysis (PCA)

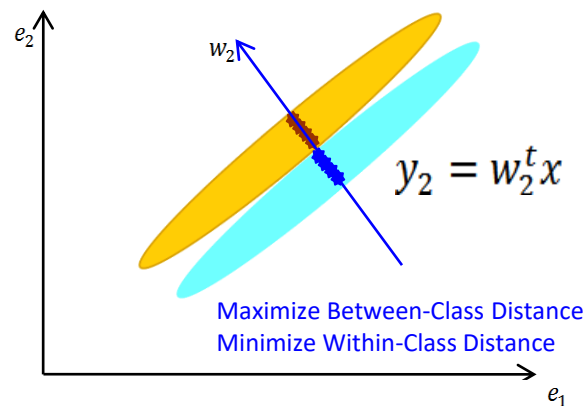
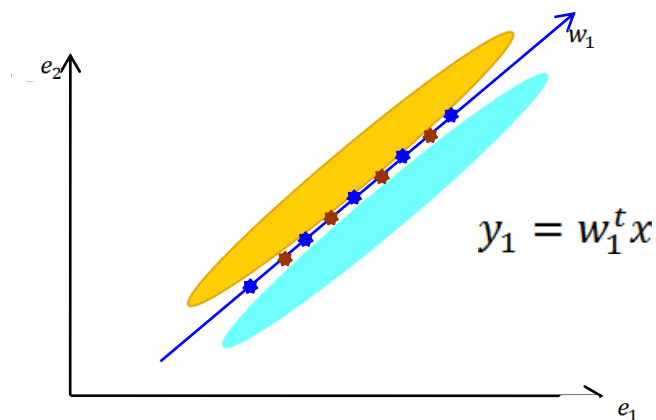
Linear Discriminant Analysis (FLDA)

Multiple Discriminant Analysis (MDA)

Simple Enhancement of PCA/LDA

Linear Discriminant Analysis: LDA

- We have n d -dimensional samples $\mathbf{x}_1, \dots, \mathbf{x}_n$, n_1 in a subset D_1 , labeled w_1 and n_2 in a subset D_2 , labeled w_2 .
- Find direction of line \mathbf{w} , that maximally separate the data.



- Let a difference between sample means be a measure of separation of projected points

Fisher Linear Discriminant cont.

- Project samples \mathbf{x}_k onto \mathbf{w} .

$$y_k = \mathbf{w}^t \mathbf{x}_k$$

- n samples y_k are divided into the subsets Y_1 and Y_2

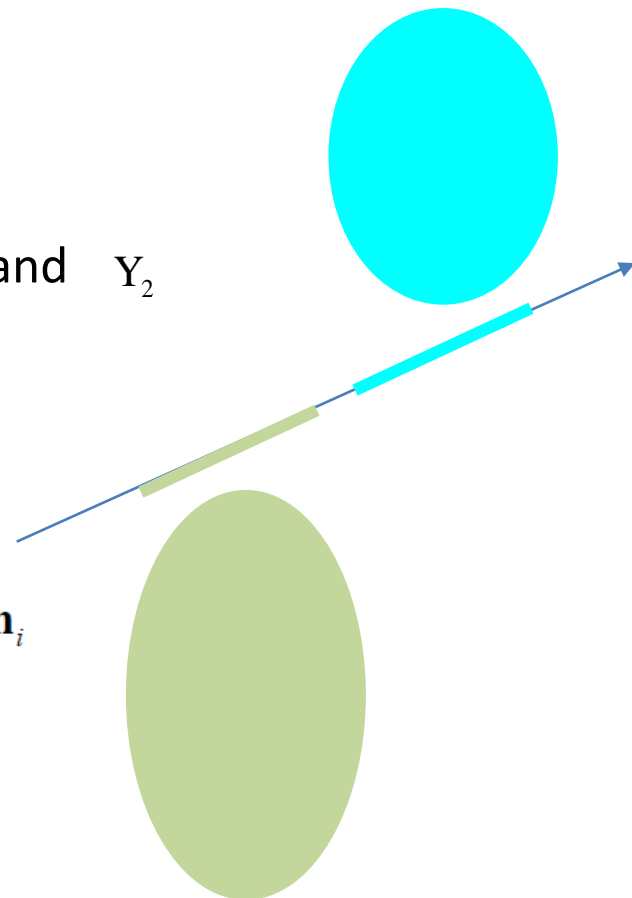
- Let \mathbf{m}_i be the sample mean
$$\mathbf{m}_i = \frac{1}{n_i} \sum_{\mathbf{x} \in D_i} \mathbf{x}$$

- The sample mean for projected points

$$\tilde{m}_i = \frac{1}{n_i} \sum_{y \in Y_i} y = \frac{1}{n_i} \sum_{\mathbf{x} \in D_i} \mathbf{w}^t \mathbf{x} = \mathbf{w}^t \mathbf{m}_i$$

- Distance between the projected means is

$$| \tilde{m}_1 - \tilde{m}_2 | = | \mathbf{w}^t (\mathbf{m}_1 - \mathbf{m}_2) |$$



Fisher Linear Discriminant cont.

- A scatter for projected samples labeled ω_i

$$\tilde{s}_i^2 = \sum_{y \in Y_i} (y - \tilde{m}_i)^2$$

$(1/n)(\tilde{s}_1^2 + \tilde{s}_2^2)$ is an **estimate of the variance** of the pooled data.

$\tilde{s}_1^2 + \tilde{s}_2^2$ is called **total within-class scatter** of the projected samples.

- The Fisher discriminant employs $\mathbf{w}^t \mathbf{x}$ for which criterion

$$J(\mathbf{w}) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{s}_1^2 + \tilde{s}_2^2}$$

is maximum

Fisher Linear Discriminant cont.

- Define scatter matrices S_i and S_w by

$$S_i = \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t$$

and

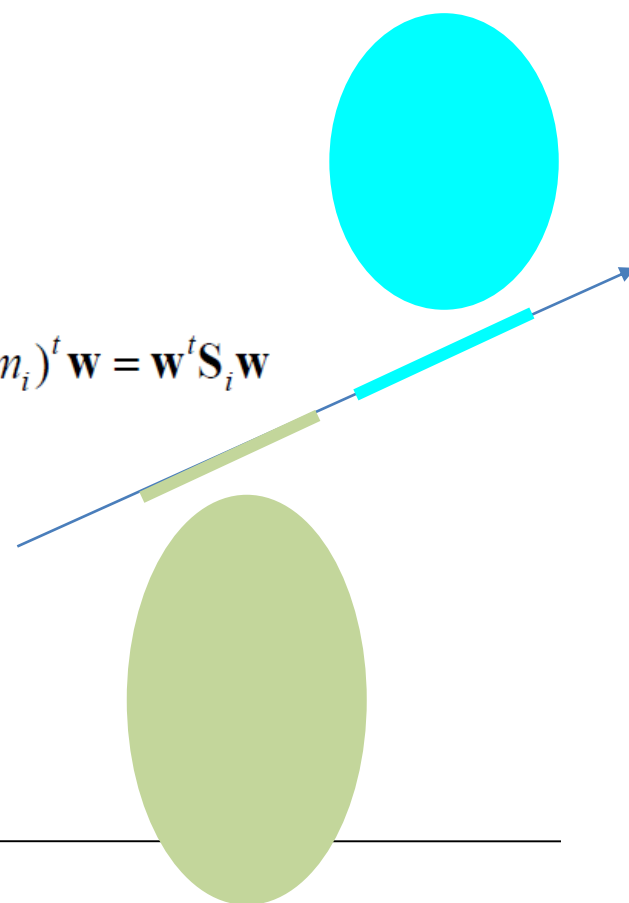
$$S_w = S_1 + S_2$$

- Then

$$\tilde{s}_i^2 = \sum_{\mathbf{x} \in D_i} (\mathbf{w}^t \mathbf{x} - \mathbf{w}^t \mathbf{m}_i)^2 = \sum_{\mathbf{x} \in D_i} \mathbf{w}^t (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t \mathbf{w} = \mathbf{w}^t S_i \mathbf{w}$$

- Thus

$$\tilde{s}_1^2 + \tilde{s}_2^2 = \mathbf{w}^t S_w \mathbf{w}$$



Fisher Linear Discriminant cont.

- Similarly,

$$(\tilde{m}_1 - \tilde{m}_2)^2 = (\mathbf{w}^t \mathbf{m}_1 - \mathbf{w}^t \mathbf{m}_2)^2 = \mathbf{w}^t (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^t \mathbf{w} = \mathbf{w}^t \mathbf{S}_B \mathbf{w}$$

\mathbf{S}_w is called **within-class scatter matrix** (proportional to sample covariance matrix)

$\mathbf{S}_B = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^t$ is called **between-class scatter matrix**.

- This gives the equivalent expression for Fisher's discriminant

$$J(\mathbf{w}) = \frac{\mathbf{w}^t \mathbf{S}_B \mathbf{w}}{\mathbf{w}^t \mathbf{S}_w \mathbf{w}}$$

- Which vector \mathbf{w} maximizes it?

$$\nabla_{\mathbf{w}} J(\mathbf{w}) = \frac{2\mathbf{S}_B \mathbf{w}}{\mathbf{w}^t \mathbf{S}_w \mathbf{w}} - \frac{\mathbf{w}^t \mathbf{S}_B \mathbf{w}}{\mathbf{w}^t \mathbf{S}_w \mathbf{w}} \frac{2\mathbf{S}_w \mathbf{w}}{\mathbf{w}^t \mathbf{S}_w \mathbf{w}} = 0$$

Fisher Linear Discriminant cont.

- Hence one gets

$$\mathbf{S}_B \mathbf{w} = \lambda \mathbf{S}_W \mathbf{w}, \quad \lambda = \frac{\mathbf{w}^t \mathbf{S}_B \mathbf{w}}{\mathbf{w}^t \mathbf{S}_W \mathbf{w}},$$

or equivalently

$$\mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w} = \lambda \mathbf{w},$$

- Since for any \mathbf{w} , $\mathbf{S}_B \mathbf{w}$ is always in the direction of $\mathbf{m}_1 - \mathbf{m}_2$:

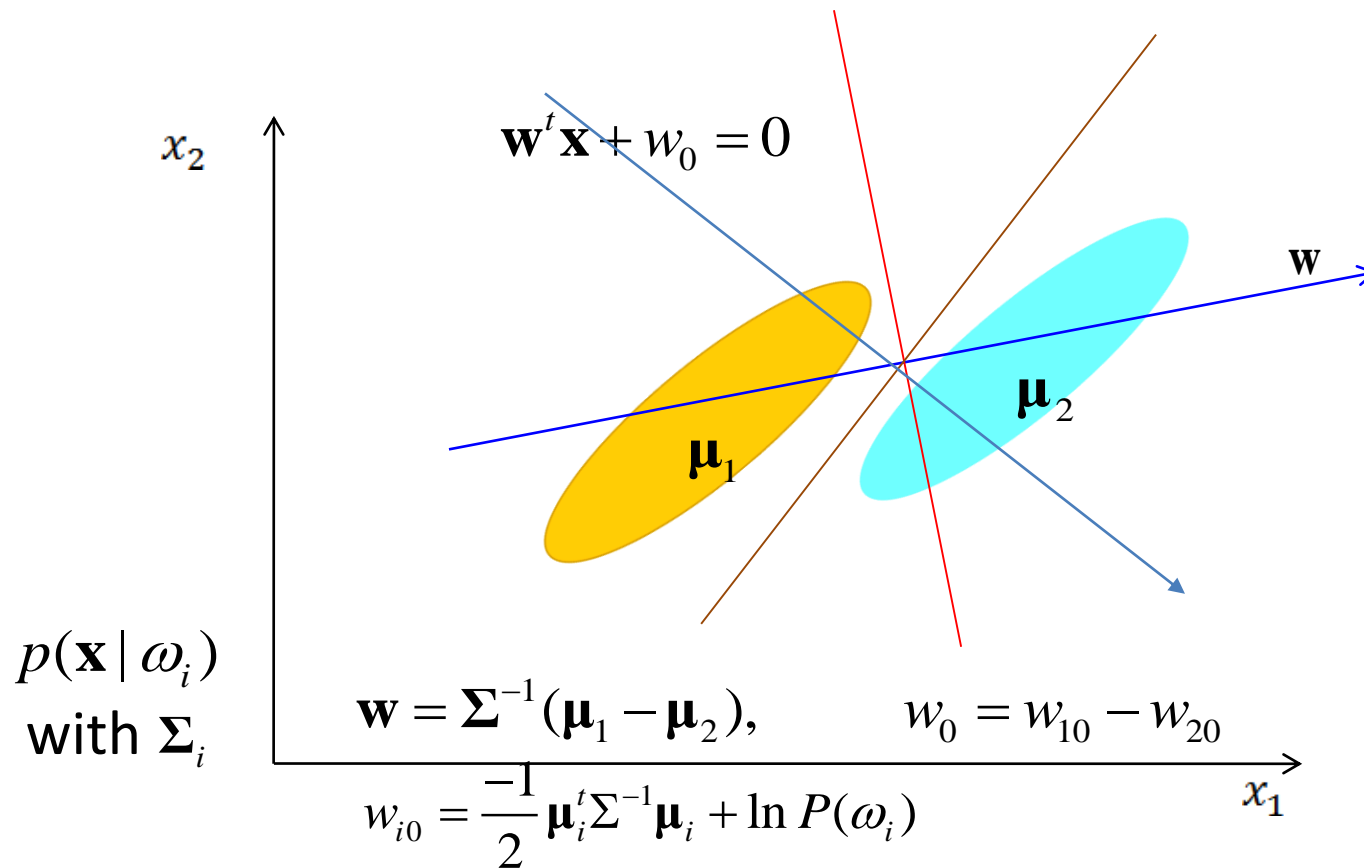
$$\mathbf{S}_B \mathbf{w} = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^t \mathbf{w} = \alpha(\mathbf{m}_1 - \mathbf{m}_2)$$

- It is not necessary to determine the eigenvalues of $\mathbf{S}_W^{-1} \mathbf{S}_B$.
- One simply gets

$$\mathbf{w} \propto \mathbf{S}_W^{-1}(\mathbf{m}_1 - \mathbf{m}_2)$$

- Scale factor for \mathbf{w} is unimportant (why?).
- FLDA is one-dimensional projection

Fisher Linear Discriminant cont.



Matrix Norm

- Induced Norm

$$\|A\|_p = \sup_{\|x\|_p=1} \|Ax\|_p \quad \|A\|_1 = \max_{1 \leq j \leq n} \sum_{i=1}^m |a_{ij}| \quad \|A\|_\infty = \max_{1 \leq i \leq m} \sum_{j=1}^n |a_{ij}|$$

- Schatten norm

$$\|A\|_p = \left(\sum_{i=1}^{\min\{m,n\}} \sigma_i^p(A) \right)^{1/p}$$

- nuclear norm

$$\|A\|_* = \text{trace}(\sqrt{A^* A}) = \sum_{i=1}^{\min\{m,n\}} \sigma_i(A)$$

- Frobenius Norm

$$\|A\|_F = \sqrt{\sum_{i=1}^m \sum_{j=1}^n |a_{ij}|^2} = \sqrt{\text{trace}(A^T A)} = \sqrt{\sum_{i=1}^{\min\{m,n\}} \sigma_i^2(A)}$$

- Spectral (maximum singular value) norm

$$\|A\|_2 = \sigma_{\max}(A) = (\lambda_{\max}(A^T A))^{1/2}$$

$$\|A\|_2 = \sup_{\|x\|_2=1} \|Ax\|_2$$

$$= \sup_{\|x\|_2=1} (x^T A^T A x)^{1/2}$$

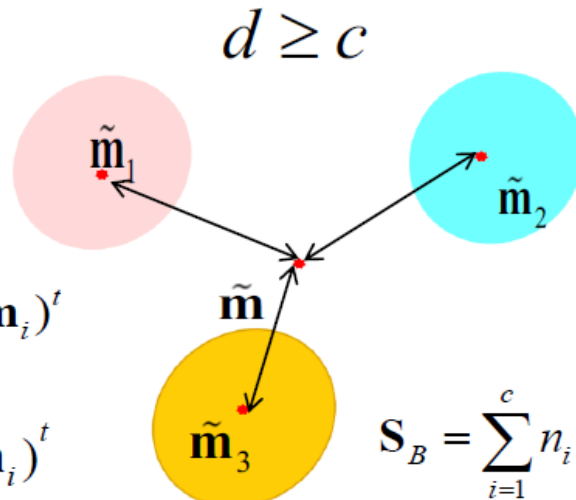
$$= \sup_{\|x\|_2=1} (x^T U^T \Lambda U x)^{1/2}$$

$$= \sup_{\|y\|_2=1} (y^T \Lambda y)^{1/2} \Leftarrow y^T y = x^T U^T U x = 1$$

$$= \sup_{\|y\|_2=1} \left(\sum y_i^2 \lambda_i \right)^{1/2}$$

$$= (\lambda_{\max}(X^T X))^{1/2}$$

Multiple Discriminant Analysis



$d \geq c$

$$S_w = \sum_{i=1}^c \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t$$

$$\tilde{S}_w = \sum_{i=1}^c \sum_{\mathbf{y} \in Y_i} (\mathbf{y} - \tilde{\mathbf{m}}_i)(\mathbf{y} - \tilde{\mathbf{m}}_i)^t$$

$$S_B = \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^t$$

$$\tilde{S}_B = \sum_{i=1}^c n_i (\tilde{\mathbf{m}}_i - \tilde{\mathbf{m}})(\tilde{\mathbf{m}}_i - \tilde{\mathbf{m}})^t$$

cf) $S_B = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^t$

Multiple Discriminant Analysis

- For the c -class problem we have $c-1$ discriminant functions.
- The projection from a d -dimensional space to a $(c-1)$ dimension is accomplished by $(c-1)$ discriminant functions (we assume $d \geq c$).

- Within-class scatter matrix is: $\mathbf{S}_w = \sum_{i=1}^c \mathbf{S}_i$

where $\mathbf{S}_i = \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t$

and $\mathbf{m}_i = \frac{1}{n_i} \sum_{\mathbf{x} \in D_i} \mathbf{x}$

- Define a total mean vector

$$\mathbf{m} = \frac{1}{n} \sum_{\mathbf{x}} \mathbf{x} = \frac{1}{n} \sum_{i=1}^c n_i \mathbf{m}_i$$

Multiple Discriminant Analysis

- And total scatter matrix $\mathbf{S}_T = \sum_{\mathbf{x}} (\mathbf{x} - \mathbf{m})(\mathbf{x} - \mathbf{m})^t$

- It can be transformed to

$$\begin{aligned}\mathbf{S}_T &= \sum_{i=1}^c \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i + \mathbf{m}_i - \mathbf{m})(\mathbf{x} - \mathbf{m}_i + \mathbf{m}_i - \mathbf{m})^t \\ &= \sum_{i=1}^c \sum_{\mathbf{x} \in D_i} (\mathbf{x} - \mathbf{m}_i)(\mathbf{x} - \mathbf{m}_i)^t + \sum_{i=1}^c \sum_{\mathbf{x} \in D_i} (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^t \\ &= \mathbf{S}_W + \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^t = \mathbf{S}_W + \mathbf{S}_B\end{aligned}$$

- The between-class scatter is:

$$\mathbf{S}_B = \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^t$$

Multiple Discriminant Analysis

- For the c -class problem we have $(c-1)$ discriminant functions. The projection from a d -dimensional space to a $(c-1)$ dimensional space is accomplished by $(c-1)$ discriminant functions:

$$y_i = \mathbf{w}_i^t \mathbf{x} \quad i = 1, \dots, (c-1)$$

- Taking d -by- $(c-1)$ \mathbf{W} matrix which columns are vectors \mathbf{w}_i , we'll get in matrix form: $\mathbf{y} = \mathbf{W}^t \mathbf{x}$
- Samples $\mathbf{x}_1, \dots, \mathbf{x}_n$ are projected to $\mathbf{y}_1, \dots, \mathbf{y}_n$.
- We define $\tilde{\mathbf{m}}_i = \frac{1}{n_i} \sum_{\mathbf{y} \in Y_i} \mathbf{y}$, $\tilde{\mathbf{m}} = \frac{1}{n} \sum_{\mathbf{y} \in Y_i} n_i \tilde{\mathbf{m}}_i$

Multiple Discriminant Analysis cont.

$$\tilde{\mathbf{S}}_W = \sum_{i=1}^c \sum_{\mathbf{y} \in Y_i} (\mathbf{y} - \tilde{\mathbf{m}}_i)(\mathbf{y} - \tilde{\mathbf{m}}_i)^t$$
$$\tilde{\mathbf{S}}_B = \sum_{i=1}^c n_i (\tilde{\mathbf{m}}_i - \tilde{\mathbf{m}})(\tilde{\mathbf{m}}_i - \tilde{\mathbf{m}})^t$$

■ It's easy to show that $\tilde{\mathbf{S}}_W = \mathbf{W}^t \mathbf{S}_W \mathbf{W}$ and $\tilde{\mathbf{S}}_B = \mathbf{W}^t \mathbf{S}_B \mathbf{W}$

■ The criterion function which should be maximized is:

$$J(\mathbf{W}) = \frac{|\tilde{\mathbf{S}}_B|}{|\tilde{\mathbf{S}}_W|} = \frac{|\mathbf{W}^t \mathbf{S}_B \mathbf{W}|}{|\mathbf{W}^t \mathbf{S}_W \mathbf{W}|} = \frac{\text{tr}(\mathbf{W}^t \mathbf{S}_B \mathbf{W})}{\text{tr}(\mathbf{W}^t \mathbf{S}_W \mathbf{W})} = \frac{\sum_i \mathbf{w}_i^t \mathbf{S}_B \mathbf{w}_i}{\sum_i \mathbf{w}_i^t \mathbf{S}_W \mathbf{w}_i}$$

■ Every column \mathbf{w}_i of \mathbf{W} we should be solution of generalized eigenvalue problem

$$\mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w}_i = \lambda_i \mathbf{w}_i$$

Multiple Discriminant Analysis cont.

- The criterion function which should be maximized is:

$$J(\mathbf{W}) = \frac{|\tilde{\mathbf{S}}_B|}{|\tilde{\mathbf{S}}_W|} = \frac{|\mathbf{W}^t \mathbf{S}_B \mathbf{W}|}{|\mathbf{W}^t \mathbf{S}_W \mathbf{W}|} = \frac{\text{tr}(\mathbf{W}^t \mathbf{S}_B \mathbf{W})}{\text{tr}(\mathbf{W}^t \mathbf{S}_W \mathbf{W})} = \frac{\sum_i \mathbf{w}_i^t \mathbf{S}_B \mathbf{w}_i}{\sum_i \mathbf{w}_i^t \mathbf{S}_W \mathbf{w}_i}$$

- Every column \mathbf{w}_i of \mathbf{W} we should be solution of generalized eigenvalue problem

$$\mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w}_i = \lambda_i \mathbf{w}_i$$

$$\mathbf{S}_B \mathbf{w}_i = \lambda_i \mathbf{S}_W \mathbf{w}_i$$

$$\mathbf{S}_B \mathbf{W} = \mathbf{S}_W \mathbf{W} \Lambda$$

$$\mathbf{W}^t \mathbf{S}_B \mathbf{W} = \mathbf{W}^t \mathbf{S}_W \mathbf{W} \Lambda$$

$$\text{tr}(\mathbf{W}^t \mathbf{S}_B \mathbf{W}) = \text{tr}(\mathbf{W}^t \mathbf{S}_W \mathbf{W}) \text{tr}(\Lambda)$$

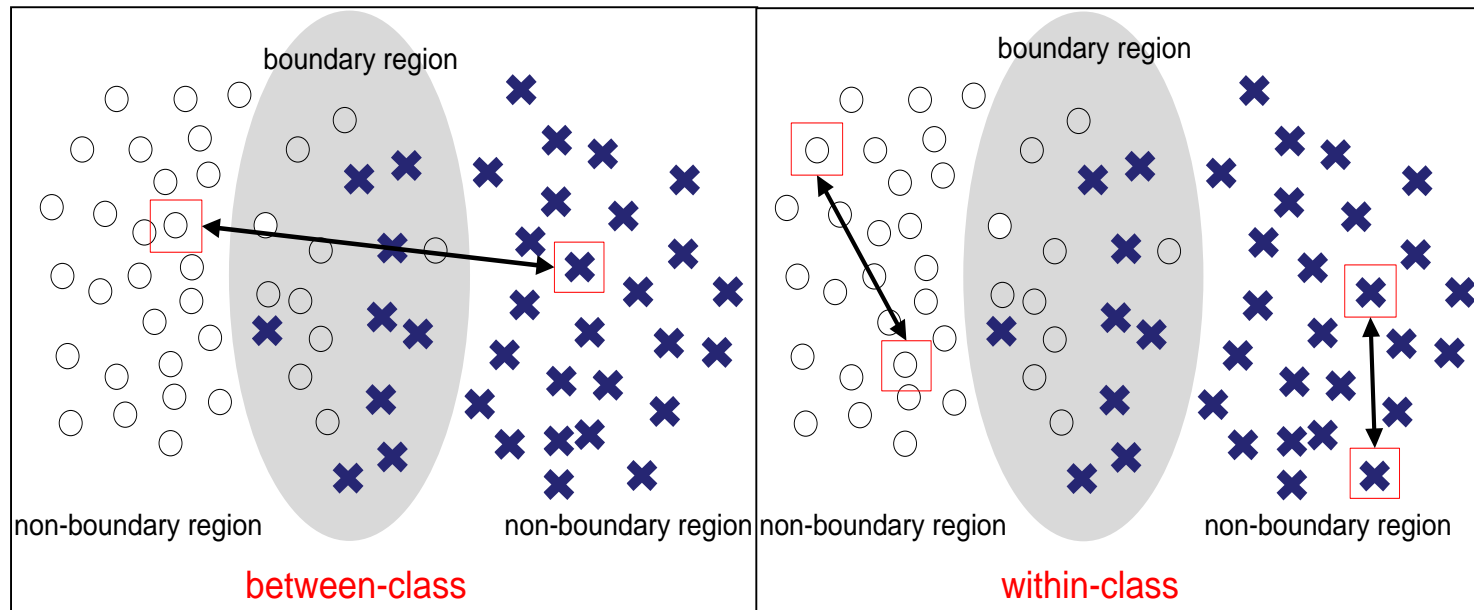
$$\text{tr}(\Lambda) = \sum_i \lambda_i = \frac{\text{tr}(\mathbf{W}^t \mathbf{S}_B \mathbf{W})}{\text{tr}(\mathbf{W}^t \mathbf{S}_W \mathbf{W})}$$

Multiple Discriminant Analysis cont.

- The MDA provides the way of reducing the dimensionality of the problem.
- The technique for finding probability density might not be feasible in the original space.
- The technique for finding probability density may work well after reducing the dimension of feature space.
- MDA may improve the separability of classes.

Simple Enhancement for PCA/LDA

- Significant pairs for **between-class scatter matrix**
 - Non-boundary patterns with the different class labels
- Significant pairs for **within-class scatter matrix**
 - Non-boundary patterns with the same class labels



Non-boundary Pattern Selection Algorithm

- Step 1. For each $\mathbf{x}_i \in \mathbf{X}$

- Find the neighborhood defined as follows.

$$Neighbors(\mathbf{x}_i, k) = N(\mathbf{x}_i, k) \cup \{\mathbf{x}_i\}$$

where $N(\mathbf{x}_i, k)$ is the set of k nearest samples to \mathbf{x}_i by L2-norm.

- Calculate voting probabilities of $Neighbors(\mathbf{x}_i, k)$ to each class j .

$$p_j(\mathbf{x}_i) = \frac{\sum_{\forall n \in Neighbors(\mathbf{x}_i, k)} I_j(n)}{k + 1}$$

where $I_j(n)$ is 1 if the class of neighbor n is j , otherwise 0.

- Calculate the neighborhood entropy of \mathbf{x}_i .

$$Neighbors_Entropy(\mathbf{x}_i, k) = \sum_{j=1}^l p_j(\mathbf{x}_i) \log_l \frac{1}{p_j(\mathbf{x}_i)}$$

- Step 2. Obtain boundary patterns $\mathbf{X}^{(B)}$ and non-boundary patterns $\mathbf{X}^{(NB)}$

$$\mathbf{X}^{(NB)} = \{\mathbf{x} | Neighbors_Entropy(\mathbf{x}, k) \leq \theta(l), \mathbf{x} \in \mathbf{X}\}$$

$$\mathbf{X}^{(B)} = \mathbf{X} - \mathbf{X}^{(NB)}$$

PCA using NPS (LDA in the same manner)

- Select non-boundary patterns via BNPS.
- Non-boundary patterns make up significant pairs.
- Emphasize the significant pairs.

Whole data

$$\mathbf{C}_X = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{m})(\mathbf{x}_i - \mathbf{m})^\top$$

$$\mathbf{W}_{PCA} = \arg \max_{\mathbf{W}^\top \mathbf{W} = \mathbf{I}} \text{tr}(\mathbf{W}^\top \mathbf{C}_X \mathbf{W})$$

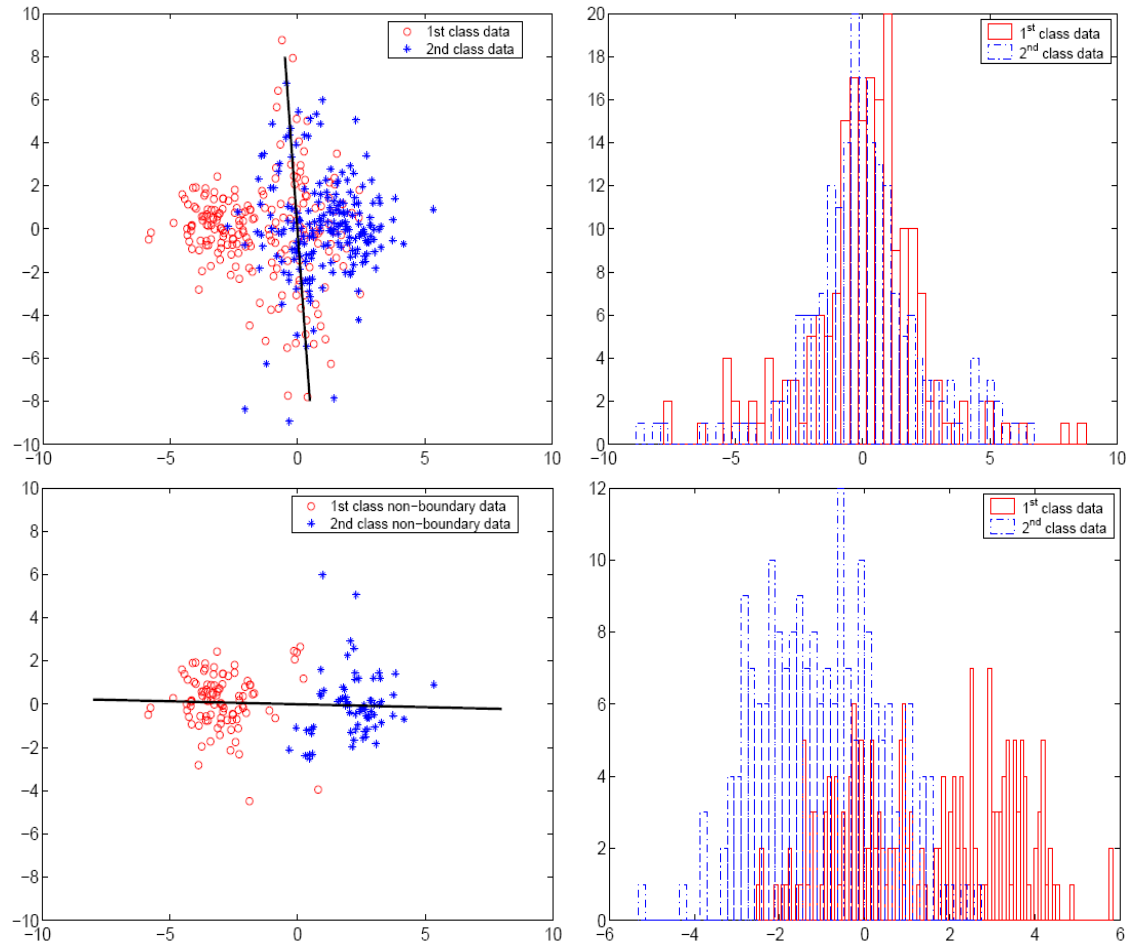


Non-boundary pattern

$$\tilde{\mathbf{C}}_X = \frac{1}{n_{NB}-1} \sum_{i=1}^l \sum_{j:y_j=i} (\mathbf{x}_j^{(NB)} - \tilde{\mathbf{m}})(\mathbf{x}_j^{(NB)} - \tilde{\mathbf{m}})^\top$$

$$\tilde{\mathbf{W}}_{PCA} = \arg \max_{\tilde{\mathbf{W}}^\top \tilde{\mathbf{W}} = \mathbf{I}} \text{tr}(\tilde{\mathbf{W}}^\top \tilde{\mathbf{C}}_X \tilde{\mathbf{W}})$$

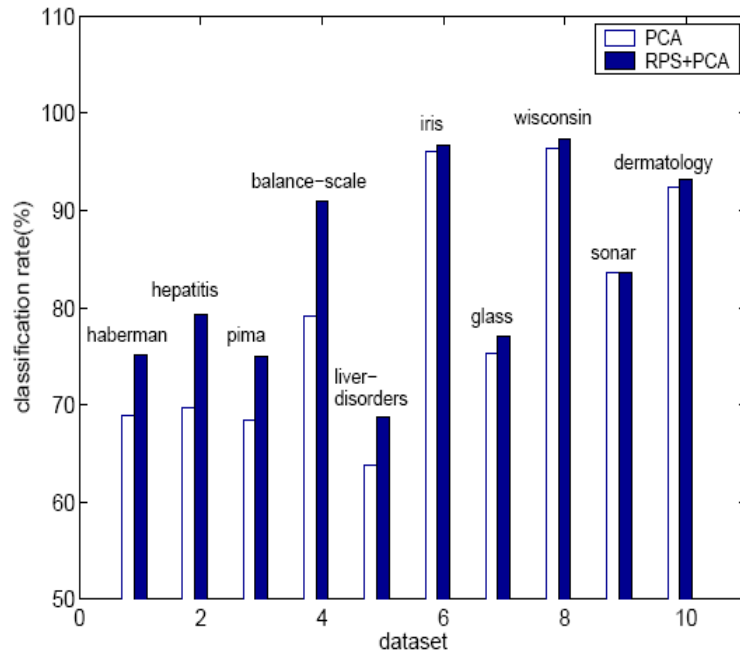
Toy Example



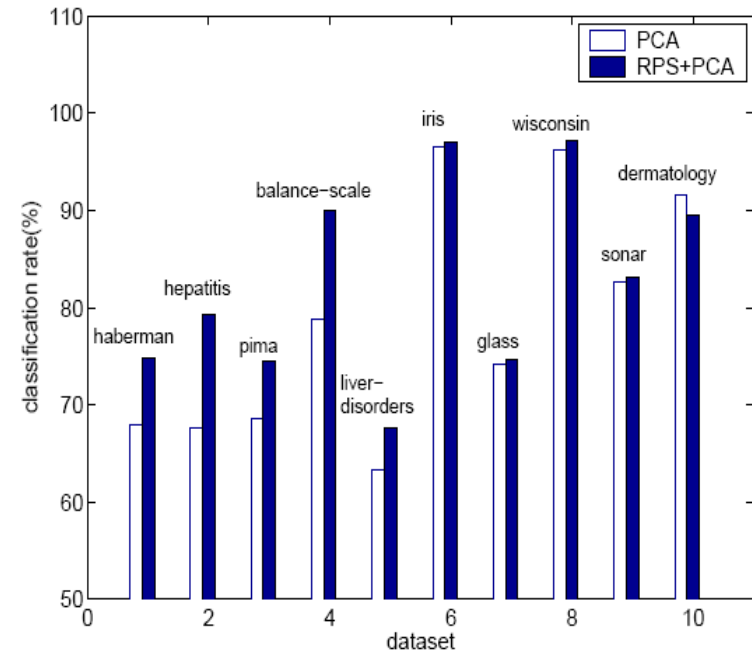
UCI Machine Learning Repository

Name	# of data	# of attributes	# of classes	Missing attributes
Haberman	306	3	2	No
Hepatitis	155	19	2	Yes
Pima	768	8	2	No
Balance-scale	625	4	3	No
Liver-disorders	345	6	2	No
Iris	150	4	3	No
Glass	214	9	6	No
Wisconsin	699	9	2	Yes
Sonar	208	60	2	No
Dermatology	366	34	6	Yes

PCA vs. NPS+PCA

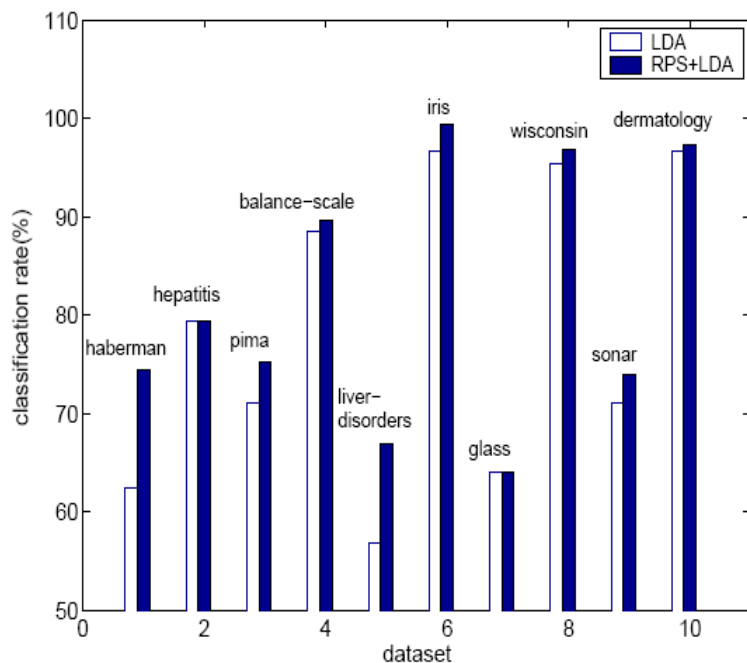


Leave-one-out
NN Classifier

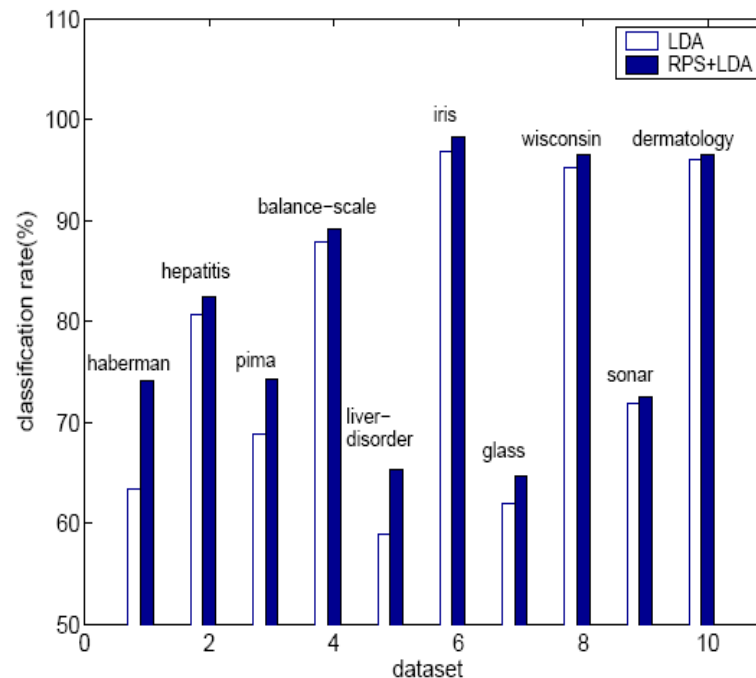


10-fold cross validation
NN Classifier

LDA vs. NPS+LDA



Leave-one-out
NN Classifier



10-fold cross validation
NN Classifier

Interim Summary

- Fisher Linear Discriminant Analysis

$$J(\mathbf{w}) = \frac{\mathbf{w}^t \mathbf{S}_B \mathbf{w}}{\mathbf{w}^t \mathbf{S}_W \mathbf{w}} \quad \mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w} = \lambda \mathbf{w}, \quad \mathbf{w} \propto \mathbf{S}_W^{-1} (\mathbf{m}_1 - \mathbf{m}_2)$$

- Multiple Discriminant Analysis

$$\mathbf{S}_B = \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m})(\mathbf{m}_i - \mathbf{m})^t \quad \mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w}_i = \lambda_i \mathbf{w}_i$$

- Simple Enhancement for PCA/LDA

$$\widetilde{\mathbf{W}}_{PCA} = \arg \max_{\widetilde{\mathbf{W}}^T \widetilde{\mathbf{W}} = \mathbf{I}} \text{tr}(\widetilde{\mathbf{W}}^T \widetilde{\mathbf{C}}_X \widetilde{\mathbf{W}})$$

$$\widetilde{\mathbf{W}}_{LDA} = \arg \max_{\widetilde{\mathbf{W}}^T \widetilde{\mathbf{W}} = \mathbf{I}} \frac{\text{tr}(\widetilde{\mathbf{W}}^T \widetilde{\mathbf{S}}^{(b)} \widetilde{\mathbf{W}})}{\text{tr}(\widetilde{\mathbf{W}}^T \widetilde{\mathbf{S}}^{(w)} \widetilde{\mathbf{W}})}$$
