# **Bayesian Network**

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# Outline

- Bayesian Networks
- Applications:
  - Traffic Pattern Analysis
  - Topic Model (Document Analysis)
- Directed Acyclic Graph
- Conditional Independence
- D-separation
- Bayesian Parameters
- Parameterized Conditional Distributions
- Multinomial, Dirichlet Distribution, Conjugate Prior
- Markov Blanket

## **Application: Traffic Pattern Analysis**

Surveillance in crowded scenes



## **Graphical Inference Model**



**Bayesian Networks** 

Directed Acyclic Graph (DAG)



$$p(a, b, c) = p(c|a, b)p(a, b) = p(c|a, b)p(b|a)p(a)$$

$$p(x_1, \ldots, x_K) = p(x_K | x_1, \ldots, x_{K-1}) \ldots p(x_2 | x_1) p(x_1)$$

## **Bayesian Networks**



$$p(x_{1}) = p(x_{1})p(x_{2})p(x_{3})p(x_{4}|x_{1}, x_{2}, x_{3}) p(x_{5}|x_{1}, x_{3})p(x_{6}|x_{4})p(x_{7}|x_{4}, x_{5})$$

**General Factorization** 

$$p(\mathbf{x}) = \prod_{k=1}^{K} p(x_k | \mathrm{pa}_k)$$

## LDA Model (Topic Modelling)





## LDA Model

Likelihood:  $p(w|z, \theta, \phi, \alpha, \beta)$ Posteriori:  $p(z, \theta, \phi|w, \alpha, \beta)$ 



•Dir(K, 
$$\alpha$$
):  $p(\mu_1, ..., \mu_K | \alpha_1, ..., \alpha_K) = \frac{\Gamma(\sum_{i=1}^K (\alpha_i))}{\prod_{i=1}^K \Gamma(\alpha_i)} \mu_1^{\alpha_1 - 1} ... \mu_K^{\alpha_K - 1}$   
•Mul(K,  $\mu$ ):  $p(x_1, ..., x_K | \mu_1, ..., \mu_K) = \frac{n!}{x_1! ... x_K!} \mu_1^{x_1} ... \mu_K^{x_K}$ 

#### Notations

*D*: the number of documents.

 $N_d$ : the number of words in *d*-th document.

K: the number of topics.

 $\alpha$ : Dirichlet prior on the per-document topic distributions.

 $\beta$ : Dirichlet prior on the per-topic word distribution.

 $\theta_d$ : topic distribution for *d*-th document.

 $\phi_k$ : word distribution for topic k.

 $z_{di}$ : the topic for the *i*-th word in d-th document.

 $w_{di}$ : the specific word.

 $\{w_{d1}, w_{d2}, ..., w_{dN_d}\}$ 

#### **Mathematical description**

Choose  $\theta_d \sim Dir(\alpha)$ . Choose  $\phi_k \sim Dir(\beta)$ . Choose a topic  $z_{ji}|\theta_d \sim Multi(\theta_d)$ . Choose a word  $w_{ji}|\phi_k, z_{di} \sim Multi(\phi_{z_{di}})$ .

*Likelihood*:  $p(w|z, \theta, \phi, \alpha, \beta)$ 

Posteriori:  $p(z, \theta, \emptyset | w, \alpha, \beta)$ 

Closed-form solution is not available

• Maximum A posteriori Probability (MAP) given observation  $w, \alpha, \beta$ 

Not Convex

$$\hat{\phi}, \hat{ heta}, \hat{z} = rg\max_{\phi, heta, z} p(\phi, heta, z | w, lpha, eta),$$

Inference of LDA Model

Bayesian Inference (Learning)  $p(\phi, heta, z | w, lpha, eta) = rac{p(\phi, heta, z, w | lpha, eta)}{p(w | lpha, eta)},$  $=rac{p(\phi, heta,z,w|lpha,eta)}{\int_{\phi}\int_{ heta}\sum\limits_{z}p(\phi, heta,z,w|lpha,eta)d heta d\phi}.$  $Z_{di}$ Likelihoods  $N_d$  $p(\phi, \theta, z, w | \alpha, \beta) = \left(\prod_{k=1}^{K} p(\phi_k | \beta)\right) \prod_{d=1}^{D} p(\theta_d \mid \alpha) \prod_{i=1}^{N_d} p(z_{di} \mid \theta_d) p(w_{di} \mid z_{di}, \phi).$ 

## **Conditional Independence**

• *a* is independent of *b* given *c* 

$$p(a|b,c) = p(a|c)$$

Equivalently

$$p(a,b|c) = p(a|b,c)p(b|c)$$
$$= p(a|c)p(b|c)$$

Notation

 $a \perp\!\!\!\perp b \mid c$ 



p(a, b, c) = p(a|c)p(b|c)p(c)

$$p(a,b) = \sum_{c} p(a|c)p(b|c)p(c)$$

 $a \not\perp b \mid \emptyset$ 

U, V, and c are independent. a = U + c, b = V + c; a, b independent?



$$p(a, b, c) = p(a|c)p(b|c)p(c)$$

$$p(a, b|c) = \frac{p(a, b, c)}{p(c)}$$

$$= p(a|c)p(b|c)$$

$$a \perp b \mid c$$

U, V, and c are independent. a = U + c, b = V + c, c = 1; a, b independent?

p(a, b, c) = p(a)p(c|a)p(b|c)

$$p(a,b) = p(a) \sum_{c} p(c|a)p(b|c) = p(a)p(b|a)$$

 $a \not\perp b \mid \emptyset$ 

p(b,c|a) = p(c|a)p(b|a,c) = p(c|a)p(b|c)



$$a \perp\!\!\!\perp b \mid c$$



p(a, b, c) = p(a)p(b)p(c|a, b)p(a, b) = p(a)p(b) $a \perp b \mid \emptyset$ 

Note: this is the opposite of Example 1, with c unobserved.

a and b are independent Bernoulli rvs. c = a + b



Note: this is the opposite of Example 1, with C observed.

a and b are independent Bernoulli rvs. c = a + b

## "Am I out of fuel?"

$$p(G = 1 | B = 1, F = 1) = 0.8$$
  

$$p(G = 1 | B = 1, F = 0) = 0.2$$
  

$$p(G = 1 | B = 0, F = 1) = 0.2$$
  

$$p(G = 1 | B = 0, F = 0) = 0.1$$

➢ G is dependent to B and F

$$p(B=1) = 0.9$$

p(F=1) = 0.9

and hence

$$p(F=0) = 0.1$$

F is independent to B



- B = Battery (0=flat, 1=fully charged)
- F = Fuel Tank (0=empty, 1=full)
- G = Fuel Gauge Reading (0=empty, 1=full)

## "Am I out of fuel?"

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$$p(B = 1) = 0.9$$
  

$$p(F = 1) = 0.9$$
  
and hence  

$$p(F = 0) = 0.1$$
  

$$p(F = 0 | G = 0) = 0$$

$$p(F = 0|G = 0) = \frac{p(G = 0|F = 0)p(F = 0)}{p(G = 0)}$$
  
\$\approx 0.257\$

Probability of an empty tank increased by observing G = 0.

## "Am I out of fuel?"

$$p(G = 1 | B = 1, F = 1) = 0.8$$
  

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$$p(G = 1 | B = 0, F = 1) = 0.2$$
  

$$p(G = 1 | B = 0, F = 0) = 0.1$$
  

$$p(B = 1) = 0.9$$
  

$$p(F = 1) = 0.9$$
  
and hence  

$$p(F = 0) = 0.1$$



$$p(F = 0|G = 0, B = 0) = \frac{p(G = 0|B = 0, F = 0)p(F = 0)}{\sum_{F \in \{0,1\}} p(G = 0|B = 0, F)p(F)}$$
  
\$\approx 0.111\$

Probability of an empty tank reduced by observing B = 0. This referred to as "explaining away".  $\succ F$  is dependent to B given G

#### Exercise

Answer the following questions for the right-hand Bayesian network.

- 1) When any random variables are not observed, show that *a* and *b* are independent to each other.
- 2) When *d* is observed, show that *a* and *b* are dependent to each other.

1) Nothing observed  

$$p(a, b, c, d) = \begin{bmatrix} & & & \\ p(a, b) = \Sigma_c \Sigma_d p(a, b, c, d) = p(a) p(b) \Sigma_c \Sigma_d \\ = p(a) p(b) \Sigma_c & & \\ p(a, b|d) = \frac{p(a, b, d)}{p(d)} = \Sigma_c \begin{cases} & & \\ \\ p(a, b|d) = \frac{p(a, b, d)}{p(d)} = \Sigma_c \begin{cases} & & \\ \\ \\ p(d) \end{bmatrix} = \frac{p(a) p(b)}{p(d)} \Sigma_c \end{cases}$$

#### Exercise

Answer the following questions for the right-hand Bayesian network.

- 1) When any random variables are not observed, show that *a* and *b* are independent to each other.
- 2) When *d* is observed, show that *a* and *b* are dependent to each other.
- 1) Nothing observed

p(a, b, c, d) = [p(a)p(b)p(c|a, b)p(d|c)]  $p(a, b) = \Sigma_c \Sigma_d p(a, b, c, d) = p(a)p(b)\Sigma_c \Sigma_d [p(c|a, b)p(d|c)]$   $= p(a)p(b)\Sigma_c [p(c|a, b)] \Sigma_d [p(d|c)]$ = [p(a)p(b)]

2) d is observed

$$p(a,b|d) = \frac{p(a,b,d)}{p(d)} = \Sigma_c \left\{ \frac{p(a,b,c,d)}{p(d)} \right\}$$
$$= \Sigma_c \left\{ \frac{p(a)p(b)p(c|a,b)p(d|c)}{p(d)} \right\} = \frac{p(a)p(b)}{p(d)} \Sigma_c \{ p(c|a,b)p(d|c) \}$$
$$= \frac{[p(a)p(b)p(d|a,b)]}{p(d)} \neq p(a|d)p(b|d)$$

#### **D-separation: Example**



#### **D**-separation

- A, B, and C are non-intersecting subsets of nodes in a directed graph.
- A path from A to B is blocked if it contains a node such that either
  - the arrows on the path meet either head-to-tail or tailto-tail at the node, and the node is in the set C, or
  - the arrows meet head-to-head at the node, and neither the node, nor any of its descendants, are in the set C.
- If all paths from A to B are blocked, A is said to be dseparated from B by C.
- If A is d-separated from B by C, the joint distribution over all variables in the graph satisfies  $A \perp B \mid C$

#### Exercise

When *B* is observed in the following Bayesian network, decide whether every path from *D* to *E* is blocked (*d*-separated) or not and determine the dependency between *D* and *E*.



#### Exercise

- a. path 1 (D $\leftarrow$ A $\rightarrow$ B $\rightarrow$ E) or (D $\rightarrow$ B $\rightarrow$ E) : Since the connection in B is head to tail and B is observed, the path 1 becomes [ ].
- b. path 2 (D  $\rightarrow$  C  $\leftarrow$  E) : Since the connection in C is head to head and C is not observed, the path 2 becomes [ ].
- c. path 3 (D  $\leftarrow$  A $\rightarrow$  E) : Since the connection in A is tail to tail and A is not observed, the path 3 becomes [ ].
- d. path 4 ( $D \rightarrow B \leftarrow A \rightarrow E$ ) : Since the connection in B is head to head and B is observed, the path  $D \rightarrow B \leftarrow A$  becomes [ ] by B. And since the connection in A is tail to tail and A is not observed, the path  $B \leftarrow A \rightarrow E$  becomes [ ]. Hence path 4 becomes [ ].
- e. Among the above 4 paths, there [exists or does not exist] at least one nonblocking path and thus D and E are [dependent or independent] to each other.



- a. path 1 (D $\leftarrow$ A $\rightarrow$ B $\rightarrow$ E) or (D $\rightarrow$ B $\rightarrow$ E) : Since the connection in B is head to tail and B is observed, the path 1 becomes [blocking].
- b. path 2 (D  $\rightarrow$  C  $\leftarrow$  E) : Since the connection in C is head to head and C is not observed, the path 2 becomes [blocking].
- c. path 3 (D  $\leftarrow$  A $\rightarrow$  E) : Since the connection in A is tail to tail and A is not observed, the path 3 becomes [non-blocking].
- d. path 4 ( $D \rightarrow B \leftarrow A \rightarrow E$ ) : Since the connection in B is head to head and B is observed, the path  $D \rightarrow B \leftarrow A$  becomes [ non-blocking ] by B. And since the connection in A is tail to tail and A is not observed, the path  $B \leftarrow A \rightarrow E$  becomes [ non-blocking ]. Hence path 4 becomes [ non-blocking ].
- e. Among the above 4 paths, there [exists or does not exist] at least one nonblocking path and thus D and E are [dependent or independent] to each other.

#### D-separation: I.I.D. Data



$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu)$$

$$p(\mathcal{D}) = \int_{-\infty}^{\infty} p(\mathcal{D}|\mu) p(\mu) \,\mathrm{d}\mu \neq \prod_{n=1}^{N} p(x_n)$$

- -

#### Discrete Variables, Multinomial

$$p(x_1, ..., x_K | \mu_1, ..., \mu_K) = \frac{n!}{x_1! ... x_K!} \mu_1^{x_1} ... \mu_K^{x_K}$$

$$p(x_{11}, \dots, x_{1K}, x_{21}, \dots, x_{2K} | \mu_{11}, \dots, \mu_{KK}) = \frac{n!}{x_{11}! \dots x_{1K}!} \frac{n!}{x_{21}! \dots x_{2K}!} \mu_{11}^{x_{11}x_{21}} \dots \mu_{KK}^{x_{1K}x_{2K}}$$



## Discrete Variables (1), Multinomial

- General joint distribution:  $K^2 1$  parameters  $p(\mathbf{x}_1, \mathbf{x}_2 | \boldsymbol{\mu}) = \prod_{k=1}^{K} \prod_{l=1}^{K} \mu_{kl}^{x_{1k}x_{2l}}$
- Independent joint distribution: 2(K 1) parameters



$$\hat{p}(\mathbf{x}_1, \mathbf{x}_2 | \boldsymbol{\mu}) = \prod_{k=1}^{K} \mu_{1k}^{x_{1k}} \prod_{l=1}^{K} \mu_{2l}^{x_{2l}}$$



 $p(x_1, x_2) = p(x_1 | x_2) p(x_2)$  K - 1 + K(K - 1)  $p(x_1, x_2) = p(x_1) p(x_2)$ K - 1 + K - 1

## Discrete Variables, Dirichlet

 The <u>posterior distributions</u> are in the same family as the <u>prior probability distribution</u>.

 $p(\mu|x) \propto p(x|\mu)p(\mu)$ 

- The prior and posterior are then called conjugate distributions, and the prior is called a conjugate prior for the likelihood function.
- Dirichlet distribution is a <u>conjugate (prior) distribution</u> to the multinomial distribution.
- Gaussian is a conjugate prior of Gaussian.



## Discrete Variables, Dirichlet

- Posteriori:  $p(\mu|x, \alpha) \propto p(x|\mu)p(\mu|\alpha)$
- $Mul(K, \mu): p(x_1, ..., x_K | \mu_1, ..., \mu_K) = \frac{n!}{x_1!...x_K!} \mu_1^{x_1} ... \mu_K^{x_K}$

• Dir(K, 
$$\alpha$$
):  $p(\mu_1, ..., \mu_K | \alpha_1, ..., \alpha_K) = \frac{\Gamma(\sum_{i=1}^K (\alpha_i - 1))}{\prod_{i=1}^K \Gamma(\alpha_i - 1)} \mu_1^{\alpha_1} ... \mu_K^{\alpha_K}$ 

- Parameters:  $\alpha_1, \dots, \alpha_K > 0$  (hyper-parameters)
- Support:  $\mu_1, ..., \mu_K \in (0,1)$  where  $\sum_{i=1}^{K} \mu_i = 1$
- Dir(K, c +  $\alpha$ ):  $p(\mu|x, \alpha) \propto p(x|\mu)p(\mu|\alpha)$ where  $c = (c_1, ..., c_K)$  is number of occurrence

• 
$$E[\mu_k] = \frac{c_k + \alpha_k}{\sum_{i=1}^K (c_i + \alpha_i)}$$



## **Discrete Variables (2)**

- General joint distribution over M variables:  $K^M 1$  parameters
- *M* -node Markov chain: K 1 + (M 1)K(K 1) parameters

 $p(x_1, x_2) = p(x_1)p(x_2|x_1)p(x_3|x_2) \dots p(x_M|x_{M-1})$ 



#### Discrete Variables: Bayesian Parameters (1)



#### Discrete Variables: Bayesian Parameters (2)



#### Parameterized Conditional Distributions



If  $x_1, \ldots, x_M$  are discrete, K-state variables,  $p(y = 1 | x_1, \ldots, x_M)$  in general has  $O(K^M)$ parameters because  $p(x_1, \ldots, x_M | y = 1)$  requires  $K^M - 1$  parameters.

The parameterized form

$$p(y=1|x_1,\ldots,x_M) = \sigma\left(w_0 + \sum_{i=1}^M w_i x_i\right) = \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x})$$

requires only M + 1 parameters (actually this can not model a probability distribution).

## The Markov Blanket

$$p(\mathbf{x}_{i}|\mathbf{x}_{\{j\neq i\}}) = \frac{p(\mathbf{x}_{1},\ldots,\mathbf{x}_{M})}{\int p(\mathbf{x}_{1},\ldots,\mathbf{x}_{M}) \, \mathrm{d}\mathbf{x}_{i}}$$
$$= \frac{\prod_{k} p(\mathbf{x}_{k}|\mathrm{pa}_{k})}{\int \prod_{k} p(\mathbf{x}_{k}|\mathrm{pa}_{k}) \, \mathrm{d}\mathbf{x}_{i}}$$
$$= \prod_{k \in MB} p(\mathbf{x}_{k}|\mathrm{pa}_{k})$$

Any factor  $p(x_k|pa_k)$  that does not have any functional dependence on  $x_i$  can be taken outside the integral over  $x_i$ , and will therefore cancel between numerator and denominator.

## LDA Model (Topic Modelling)



#### Topic proportions and Documents assignments Seeking Life's Bare (Genetic) Necessities COLD SPRING HARBOR, NEW YORK-"are not all that far apart," especially in How many genes does an organism need to comparison to the 75,000 genes in the huan genome, notes Siv Andersson of survive! Last week at the genome meeting University in Sweden, the arrived at here,8 two genome researchers with radically different approaches presented complemen-800 number. But coming up with a c tary views of the basic genes needed for life sus answer may be more than just a One research team, using computer analynumbers some particularly, ses to compare known penomes, concluded more genomes are completely sequenced. "It may be a way of organil that today's organisms can be sustained with just 250 genes, and that the earliest life forms any newly sequenced genome," explains required a mere 128 genes. The Arcady Mushegian, a computational molecular biologist at the National Center other researcher mapped genes in a simple parasite and estifor Biotechnology Information (NCBI) mated that for this organism, in Bethesda, Maryland, Comparing genome 1703 gete 800 genes are plenty to do the Redundant and job-but that anything short Genes -122 perces of 100 wouldn't be enough. Although the numbers don't match precisely, those predictions gerautre 459 paries \* Genome Mapping and Sequenc- $\theta_d$ ing, Cold Spring Harbor, New York, Stripping down. Computer analysis yields an esti-May 8 to 12. mate of the minimum modern and ancient genomes. SCIENCE • VOL. 272 • 24 MAY 1996 $Z_{di}$ $\phi_k$ Nd

## **Application: Traffic Pattern Analysis**

Surveillance in crowded scenes



## Hidden Markov Model



 $p(x_1, \dots, x_T) = p(x_1)p(x_2|x_1) p(x_3|x_2) \dots \dots p(x_T|x_{T-1})$ 

## Variational Auto-encoder (VAE)



## **Interim Summary**

- Bayesian Networks
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- D-separation
- Bayesian Parameters
- Parameterized Conditional Distributions
- Multinomial, Dirichlet Distribution, Conjugate Prior
- Markov Blanket
- Applications:
  - Topic Model (Document Analysis, Traffic Pattern Analysis)
  - Hidden Markov Model
  - Generative Models